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WORKING PAPER SERIES

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Massimiliano Marcellino

Working Paper n. 211

March 2002

IGIER – Università Bocconi, Via Salasco 5, 20136 Milano –Italy
<http://www.igier.uni-bocconi.it>

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Instability and non-linearity in the EMU *

Massimiliano Marcellino **

IEP-Università Bocconi, IGER and CEPR

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Abstract

In this paper we evaluate the relative performance of linear, non-linear and time-varying models for about 500 macroeconomic variables for the countries in the Euro area, using a real-time forecasting methodology. It turns out that linear models work well for about 35% of the series under analysis, time-varying models for another 35% and non-linear models for the remaining 30% of the series. The gains in forecasting accuracy from the choice of the best model can be substantial, in particular for longer forecast horizons. These results emerge from a detailed disaggregated analysis, while they are hidden when an average loss function is used. To explore in more detail the issue of parameter instability, we then apply a battery of tests, detecting non-constancy in about 20-30% of the time series. For these variables the forecasting performance of the time-varying and non-linear models further improves, with larger gains for a larger fraction of the series. Finally, we evaluate whether non-linear models perform better for three key macroeconomic variables: industrial production, inflation and unemployment. It turns out that this is often the case. Hence, overall, our results indicate that there is a substantial amount of instability and non-linearity in the EMU, and suggest that it can be worth going beyond linear models for several EMU macroeconomic variables.

J.E.L. Classification: C2, C53, E30

Keywords: Instability, Non-linearity, Time-Varying models, Non-linear models, European Monetary Union

* I am grateful to Jim Stock and Mark Watson for several conversations on the topics addressed in this paper, and for making available the GAUSS programs that form the basis of the code used in this paper. Seminar participants at the ECB also provided useful comments. The usual disclaimers apply.

** Correspondence to: IGER, Università Bocconi, Via Salasco 3/5, 20136, Milan, Italy.
E-mail: massimiliano.marcellino@uni-bocconi.it Phone: +39-02-5836-3327 Fax: +39-02-5836-3302

1. Introduction

The many social, economic and political changes that occurred in the European countries since the early '80s can be expected to make modelling macroeconomic variables with constant parameter linear models particularly difficult. In this context, time-varying and non-linear models should have a comparative advantage, even though their estimation is complicated by the short time span usually available. A comparison of the in sample goodness of fit of linear and non-linear models would be biased in favor of the latter, because of their extensive parameterization. Hence, the relative performance of linear and non-linear models is better evaluated out of sample, with a real time forecasting exercise.

In this paper we conduct such a comparison, using a very large dataset of about 500 macroeconomic variables for the countries in the Euro area. A similar analysis is performed by Stock and Watson (1999) for the US. Here we have a larger dataset, for several countries, but for a shorter sample. Moreover, we consider more forecasting models, and evaluate the role of instability in explaining the results.

We compare three main forecasting methods. The linear method, which includes autoregressive (AR) models, exponential smoothing and random walk models. The time-varying method, which includes time-varying AR models and smooth transition AR models. The non-linear method, that includes artificial neural network models. Within each method we consider several alternative specifications, for a total of 58 models. We focus on three forecast horizons: one, three and six month a-head. Longer horizons are not worth analyzing because of the rather short sample available.

The competing forecasts are compared on the basis of three measures, in increasing level of disaggregation. First, we compute the average value over all variables of several loss functions, including the common mean absolute and mean square forecast error. Second, we rank the models on the basis of the percentage of variables for which they are among the top-N models, for several values of N. Finally, for each forecasting model we compute the empirical distribution function (over variables) of its mean square forecast error, relative to a benchmark AR model, and report selected percentiles of this distribution.

It turns out that linear models work well for about 35% of the series under analysis, time-varying models for another 35% and non-linear models for the remaining 30% of

the series. The gains in forecasting accuracy from the choice of the best model can be substantial, in particular for longer forecast horizons. These results emerge from a detailed disaggregated analysis, while they are hidden when an average loss function is used.

To explore whether the results are driven by extensive parameter instability, we then apply a battery of stability tests, including statistics for constancy of the parameters versus the alternative of random walk or stationary random coefficients, and recursive F-tests for parameter changes at unknown dates. We detect non-constancy in about 20-40% of the time-series, though the percentage decreases to 10-20% when using bootstrapped critical values for the test statistics. The forecasting performance of the time-varying and non-linear models for the unstable variables further improves, with larger gains for a larger fraction of the series.

Finally, we evaluate whether non-linear models perform better for three key macroeconomic variables: industrial production, inflation and unemployment. It turns out that this is often the case.

The structure of the paper is the following. The dataset is illustrated in Section 2. Section 3 describes the forecasting models and the evaluation criteria. The results of the forecast comparison are presented in Section 4. The issue of instability is analyzed in Section 5. Section 6 presents detailed results for industrial production, unemployment and inflation. Section 7 summarizes and offers some concluding remarks.

2. The data

The dataset we use is taken from Marcellino, Stock and Watson (2000,2001), to whom we refer for additional information. It includes the OECD main economic indicators, monthly, for the period 1982:1-1997:8, for the 11 countries originally in the EMU in the year 2000. The dataset and the sample range is chosen in order to have rather homogenous variables over countries, for a long enough comparable time span. Overall, there are 480 series.

In particular, for each country there are output variables (industrial production and sales, disaggregated by main sectors); labour market variables (employment,

unemployment, wages and unit labour costs); prices (consumer and producer, disaggregated by type of goods); monetary aggregates, interest rates (different maturities), stock prices; exchange rates (effective and nominal); imports, exports and net trade; and other miscellaneous series. A complete list of the variables is reported in the Appendix.

3 Forecasting methods

The formulation of a generic forecasting model is

$$y_{t+h}^h = f(Z_t; \mathbf{q}_{ht}) + \mathbf{e}_{t+h}, \quad (1)$$

where y_t is the variable being forecast, h indicates the forecast horizon, Z_t is a vector of predictor variables, \mathbf{e}_t is an error term, and \mathbf{q}_h is a vector of parameters, possibly evolving over time. We introduce a distinction between forecasting methods and forecasting models. Forecasting methods differ for the choice of the functional form of the relationship between y_{t+h}^h and Z_t , f . Within each method, different models are determined by the choice of the regressors Z_t and the stationarity transformation applied to y_t .

The h -step forecast is

$$\hat{y}_{t+h}^h = f(Z_t; \hat{\mathbf{q}}_{ht}), \quad (2)$$

with associated forecast error

$$e_{t+h} = y_{t+h}^h - \hat{y}_{t+h}^h. \quad (3)$$

When y_t is treated as stationary, it is $y_{t+h}^h = y_{t+h}$, while if y_t is I(1) then $y_{t+h}^h = y_{t+h} - y_t$. We present results for both cases. Moreover, we also consider a pre-test forecast where the decision on the stationarity of y_t is based on a unit root test, which often improves the forecasting performance, see e.g. Diebold and Kilian (2000). In particular, we use the Elliott, Rothenberg and Stock (1996) DF-GLS statistics, which performed best in the simulation experiments in Stock (1996). Note that $e_{t+h} = y_{t+h} - \hat{y}_{t+h}$, independently of

whether y_t is treated as stationary or not, so that forecast errors from the three different cases (stationary, I(1) and pre-test) are directly comparable.

Because of the short sample period available, the forecast horizons we consider are 1, 3 and 6 months. When h is larger than one, the " h -step ahead projection" approach in (1), also called dynamic estimation (e.g. Clements and Hendry (1996)), differs from the standard approach of estimating a one-step ahead model, then iterating that model forward to obtain h -step ahead predictions. The h -step ahead projection approach has two main advantages in this context. First, the potential impact of specification error in the one-step ahead model can be reduced by using the same horizon for estimation as for forecasting. Second, we need not resort to simulation methods to obtain forecasts from non-linear models. The resulting forecasts could be slightly less efficient, see e.g. Granger and Terasvirta (1993, Ch.8), but the computational savings in our real time exercise with many series are substantial.

A few forecast errors from the non-linear and time-varying methods are very large. This is due to problems in the estimation of these models, because of multiple local optima and the short estimation sample available (1982:1-1993:12). In order not to bias the comparison against these methods, we automatically trim the forecasts. In particular, when the absolute value of a forecasted change is larger than any previously observed change, a no change forecast is used.

Let us now list the methods and models we compare, and briefly discuss their main characteristics and estimation issues. More details can be found in Stock and Watson (1996, 2000).

Linear methods

Autoregression (AR). Box and Jenkins (1970) popularized the use of these models for forecasting economic variables, and they have performed rather well in forecast comparison exercises, see e.g. Meese and Geweke (1984), or Marcellino, Stock and Watson (2001) for the Euro area. The f function in (1) is linear, and Z_t includes lags of the y variable and a deterministic component. The latter can be either a constant or also a linear trend. The lag length is either fixed at 4, or it is chosen by AIC or BIC with a

maximum of 6 lags. Recalling that the y_t variable can be treated as stationary, I(1), or pre-tested for unit roots, overall we have 18 models in this class.

Exponential smoothing (ES). Exponential smoothing is a rather simple forecasting technique that can be shown to be optimal in the mean square forecast error sense only when the underlying process follows a particular ARMA structure, see e.g. Granger and Newbold (1986, Ch.5). Yet, Makridakis et al. (1982) found this method to perform rather well in practice in a forecast comparison exercise. We consider both single and double exponential smoothing, which are usually adopted for, respectively, stationary and trending series. Estimation of the parameters is conducted by means of (recursive) non-linear least squares (see e.g. Tiao and Xu (1993)). The third model in this class is given by a combination of the single and double models, based on the outcome of the unit root test.

No change. This simple forecast is based on a random walk model, so that it is $\hat{y}_{t+h} = y_t$. Notwithstanding its simplicity, in a few cases it was found to outperform even forecasts from large-scale structural models, see e.g. Artis and Marcellino (2001).

Time-varying methods

Time-varying autoregression (TVAR). In this case the parameters of the AR models evolve according to the following multivariate random walk model (see e.g. Nyblom (1989)):

$$\mathbf{q}_{ht} = \mathbf{q}_{ht-1} + \mathbf{u}_{ht}, \quad \mathbf{u}_{ht} \sim iid(0, \mathbf{I}^2 \mathbf{s}^2 Q), \quad (4)$$

where \mathbf{s}^2 is the variance of the error term \mathbf{e} in (1), $Q = (E(Z_t Z_t'))^{-1}$, and we inspect several values of \mathbf{I} : 0 (no evolution), 0.0025, 0.005, 0.0075, 0.01, 0.015, or 0.020. We consider first a specification with a constant, 3 lags and $\mathbf{I} = 0.005$, and then we allow for selection of the number of lags (1,3,6) jointly with the value of \mathbf{I} by either AIC or BIC. In each case, y_t can be either stationary, or I(1) or pre-tested, so that we have a total of 9 TVAR models. The models are estimated by the Kalman filter.

Logistic smooth transition autoregression (LSTAR). The generic model can be written as

$$y_{t+h}^h = \mathbf{a}' \mathbf{z}_t + d_t \mathbf{b}' \mathbf{z}_t + \mathbf{e}_{t+h}, \quad (5)$$

where $d_t = 1/(1 + \exp(\mathbf{g}_0 + \mathbf{g}_1 \mathbf{z}_t))$, and $\mathbf{z}_t = (1, y_t, y_{t-1}, \dots, y_{t-p+1})$ if y_t is treated as stationary or $\mathbf{z}_t = (1, \Delta y_t, \Delta y_{t-1}, \dots, \Delta y_{t-p+1})$ if y_t is I(1). The smoothing parameters \mathbf{g}_1 regulate the shape of parameter change over time. When $\mathbf{g}_1 = 0$ the model becomes linear, while for large values of \mathbf{g}_1 the model tends to a self-exciting threshold model, see e.g. Granger and Terasvirta (1993), Terasvirta (1998) for details. For models specified in levels we consider the following choices for the threshold variable in d_t : $\mathbf{z}_t = y_t$, $\mathbf{z}_t = y_{t-2}$, $\mathbf{z}_t = y_{t-5}$, $\mathbf{z}_t = y_t - y_{t-6}$, $\mathbf{z}_t = y_t - y_{t-12}$. For differenced variables, it can be $\mathbf{z}_t = \Delta y_t$, $\mathbf{z}_t = \Delta y_{t-2}$, $\mathbf{z}_t = \Delta y_{t-5}$, $\mathbf{z}_t = y_t - y_{t-6}$, $\mathbf{z}_t = y_t - y_{t-12}$. In each case the lag length of the model was either 1 or 3 or 6. We report results for the following models: 3 lags and $\mathbf{z}_t = y_t$ (or $\mathbf{z}_t = \Delta y_t$ for the I(1) case); 3 lags and $\mathbf{z}_t = y_t - y_{t-6}$; AIC or BIC selection of both the number of lags and the specification of \mathbf{z}_t . In each case, y_t can be either stationary, or I(1) or pre-tested, so that overall there are 12 LSTAR models. Estimation is carried out by (recursive) non-linear least squares, using an optimizer developed by Stock and Watson (1999).

Non-linear methods

Artificial neural network (ANN). Artificial neural networks can provide a valid approximation to the generating mechanism of a vast class of non-linear processes, see e.g. Hornik, Stinchcombe and White (1989), and Swanson and White (1997) for their use as forecasting devices. The so called single layer feedforward neural network model with n_1 hidden units (and a linear component) is specified as:

$$y_{t+h}^h = \mathbf{b}_0' \mathbf{z}_t + \sum_{i=1}^{n_1} \mathbf{g}_{1i} g(\mathbf{b}_{1i}' \mathbf{z}_t) + \mathbf{e}_{t+h}, \quad (6)$$

where $g(x)$ is the logistic function, $g(x) = 1/(1 + e^{-x})$. Note that when $n_1=1$ the model can be interpreted as a logistic smooth transition autoregression, with the parameter evolution being determined by the linear combination of variables $\mathbf{b}_{11}' \mathbf{z}_t$. A more complex model is the double layer feedforward neural network with n_1 and n_2 hidden units:

$$y_{t+h}^h = \mathbf{b}_0' \mathbf{z}_t + \sum_{j=1}^{n_2} \mathbf{g}_{2j} g\left(\sum_{i=1}^{n_1} \mathbf{b}_{2ji} g(\mathbf{b}_{1i}' \mathbf{z}_t)\right) + \mathbf{e}_{t+h}. \quad (7)$$

We report results for the following specifications: $n_1=2, n_2=0, p=3$ (recall that p is number of lags in \mathbf{z}_t); $n_1=2, n_2=1, p=3$; $n_1=2, n_2=2, p=3$; AIC or BIC selection with $n_1=(1,2,3), n_2=(1,2 \text{ with } n_1=2), p=(1,3)$. For each case y_t can be either stationary, or I(1) or pre-tested, which yields a total of 15 ANN models. The models are estimated by (recursive) non-linear least squares, using an algorithm developed by Stock and Watson (1999).

Overall, there are 58 models in the forecast comparison exercise, 22 belong to the linear class, 21 are time-varying, and 15 are non-linear. They are summarized in Table 1. To mimic real time situations, for each variable, method and model the unit-root tests, estimation and model selection are repeated each month over the forecasting period, which is 1994:1-1997:8.

4. Forecast Evaluation

We now have to evaluate the relative forecasting performance of the $M=58$ models for the $N=480$ variables in the dataset. The starting point is the choice of a loss function.

For variable n and forecasting method m , we define the loss function as

$$Loss_{n,m}^h = \frac{1}{T-h} \sum_{t=1}^{T-h} |e_{t+h,n,m}|^r, \quad (8)$$

where e_{t+h} is the h -step ahead forecast error, and r can be equal to 1, 1.5, 2, 2.5 or 3. The values $r=1$ and $r=2$ correspond to the familiar choices of, respectively, the mean absolute and the mean square forecast error as the loss function.

In order to compare the loss over the whole set of variables, we adopt the following loss function for method m :

$$Loss_m^h = \frac{1}{N} \sum_{n=1}^N \frac{Loss_{n,m}^h}{Loss_{n,1}^h}, \quad (9)$$

namely, a weighted average of the loss for each variable, with weights given by the inverse of the loss of a benchmark forecast, which makes the magnitude of the losses

comparable across variables. As a benchmark, we adopt throughout an AR model with 4 lags and a constant, specified in levels.

In Table 2 we report the ranking of the models, for different values of \mathbf{r} . The results are quite clear cut, and three main comments are in order. First, the best model for any value of \mathbf{r} is an AR, with 4 lags when $h=1$ or 3 (our benchmark), and with AIC lag selection for $h=6$. The AR(4) ranks second for $h=6$, and the AR(AIC) ranks second for $h=1,3$. In both cases, it is better not to impose the presence of a unit root and not to include a linear trend in the model. AR models with unit roots imposed or pre-tested for appear in the ranking at, respectively, the third and fourth place for some values of \mathbf{r} and h .

Second, time-varying models perform better than non-linear models, and the best time-varying models are the TVARs. In particular, when $\mathbf{r}=2$, the TVAR with 3 lags, constant and no unit root imposed ranks 4th for $h=6$, 10th for $h=3$ and 17th for $h=1$. Several other TVAR models are ranked 15th or lower, while the only LSTAR model in the top-15 has 3 lags, no unit root, and the transition variable is $\mathbf{z}_t = y_t - y_{t-6}$.

Third, the performance of the neural network models is very poor, the best ranked model is only 35th. To make sure that this result is not driven by estimation problems due to the small sample size, we have repeated the estimation and forecasting exercises several times with different starting values and a very large number of iterations in the optimization routines, but the figures did not change. Moreover, Stock and Watson (1999) got similar results with longer series for the US.¹

The ranking in Table 2 is based on the loss function in (9), which is an average over all the variables, so that the TV and non-linear models could still be valuable for some series. To evaluate whether this is the case, we adopt the loss function in (8) with $\mathbf{r}=2$, i.e. the mean square forecast error (msfe), and we find the best forecasting model for each

¹ Note that since we are comparing the average loss over about 500 variables, it is virtually impossible to provide standard errors for the measure in equation (9), and hence tests for whether the average loss between two models is statistically different from zero. Yet, the ranking in Table 2 definitely favors linear models.

variable. In Table 3a we report the fraction of series for which a given method is the best (results for each model are available upon request).

The picture is now rather different. The best methods for the largest proportion of series are ANN, which leads to the lowest msfe for 27% of the 480 series when $h=1$ and for 32% when $h=6$, and LSTAR, with 28% of the lowest msfe for $h=3$. The AR method ranks third, with values of 23% for $h=1$ and 6, and of 20% for $h=3$. Even though Stock and Watson (1999) do not explicitly mention it, a similar pattern emerges also for the US, see their Tables 3 and 4.

The second panel of Table 3a shows that if we aggregate the models into linear, time-varying and non-linear methods, there is a substantial equivalence of the three approaches in terms of the fraction of best forecasting models, with only slightly lower values for the non-linear approach.

In Table 3b we provide additional results for the models that perform best for the highest fraction of series in each class. In particular, we compute the fraction of series for which these models are the best or among the top 5, 10, 15 and 20. The resulting figures are of comparable magnitude across models, with slightly lower values for the no-change forecast.

To explain the mismatching ranking of Tables 2 and 3 we have to take a more disaggregate approach. First, for each variable we compute the relative msfe (rmsfe) of each forecasting model with respect to the benchmark AR(4), so that an rmsfe higher than one indicates that the method under analysis is worse than the benchmark. In formulae, the rmsfe of model j for variable m is:

$$rmsfe^h_{j-AR4,m} = \left(\sum_{t=1}^{T-h} e^2_{j,t+h,m} \right) / \left(\sum_{t=1}^{T-h} e^2_{AR4,t+h,m} \right) \quad (10)$$

Then, for each model, we calculate the empirical distribution of the rmsfe over the variables. In Table 4 we report the mean of the distribution and some percentiles for selected models (the best in Table 2 and those in Table 3b, results for all models are available upon request).

The ranking in Table 2 for $\mathbf{r}=2$ is based on the mean of this distribution, while that in Table 3 in practice is based on the lower percentiles. Hence, we expect the ARFC0a (and the benchmark) to have a lower average rmsfe and a more concentrated

distribution, while the ANN models, and more generally the models in Table 3b, should present higher dispersion and good performance in the lower tail of the distribution.

The figures in Table 4 confirm our expectations. In particular, values of the rmsfe in the lower 10% tail of the distribution are substantially smaller than one for ANN, but also for LSTAR models, while some values in the upper 10% tail are very large. This could suggest to use the median rather than the mean of the distribution to construct the ranking in Table 2. Yet, from the 6th column of Table 4 the median is also in general higher for non-linear than for linear models. It is important to take into consideration the whole distribution.

It is also worth noting that the AR in levels with a constant and AIC lag length selection (ARFC0a) has a highly concentrated distribution over variables, most rmsfe are in the range 0.85 to 1.19 for $h=1$. The corresponding values for the ANN model ANF0b are 0.66 to 2.15, and 0.68 to 2.04 for the LSTAR model LS0063 (these are the best models in their class from Table 3b). Moreover, for these models the dispersion of the distribution increases much more rapidly with the forecast horizon, h , than for the AR model.

These characteristics imply that the AR model has a much more stable performance over variables and forecast horizons than time-varying and non-linear models, but the latter can yield substantial gains for some variables and forecast horizons. In Section 6 we will evaluate whether this is the case for some key macroeconomic variables, while in the next Section we analyze in more detail the issue of instability.

5. Measuring the extent of instability

Time-varying models work better than linear specifications for about one third of the series under analysis. Moreover, the success of the non-linear models can be also partly due to instability. In this section we try to measure the extent of instability in the Monetary Union by applying a battery of parameter constancy tests to all the 480 macroeconomic variables. We then repeat the forecasting exercise for the subset of unstable series, and verify whether the time-varying and non-linear models do perform better in this case.

5.1 Instability Tests

Following Stock and Watson (1996), who present a detailed analysis of instability for the US, we consider three different types of statistics.

First, tests for constant versus randomly time-varying coefficients. This set includes Nyblom's (1989, NY) locally most powerful test against the alternative of random walk coefficients ($\lambda=0$ versus $\lambda >0$ in equation (4)), and a Breusch and Pagan (1979, BP) Lagrange multiplier test against the alternative of iid random coefficients with constant mean and variance.

Second, tests based on functions of the cumulative sum of OLS residuals from equation (1), see Ploberger and Kramer (1992). We consider the supremum of the cumulative sum (KP1), and its mean square (KP2).

Third, F-tests for constancy of the parameters against the alternative of a single break at an unknown date. The tests are computed recursively for a range of dates, say $[t_0, \dots, t_1]$, where t_0 and t_1 are selected in order to discard the first and last 15% of the sample. Three functions of the resulting sequence of statistics are considered: the supremum (Quandt (1960, QLR)); the mean (Hansen (1992), Andrews and Ploberger (1994), MLR); and the so called average exponential (Andrews and Ploberger (1994), ALR).

Stationarity transformations, i.e. logarithms and differencing, are applied to all series when needed (a detailed list is available upon request), and all series are represented as an AR process in levels, with 3 lags and a constant. In Table 5a we report results for the stability tests, using different significance levels and asymptotic critical values from the papers listed above. More precisely, we report the percentage of series for which the null hypothesis of stability is rejected.

Using a 10% significance level, the figures are in the range 20-40%, with lower values from the KP1 and KP2 tests, and higher values from the recursive F-tests. There is no clear-cut pattern of rejection of stability among series and countries (detailed results for each variable are available upon request, see also the Data Appendix and Section 6), even though price series and monetary aggregates appear often to be unstable.

In Table 5b we investigate whether the detected amount of instability is robust to three extensions. First, instead of using a fixed lag length for all the series, we allow for BIC selection for each variable. Second, because of the short sample, the finite sample distribution of the tests could be different from its asymptotic counterpart. To address this issue, we have generated finite sample critical values by bootstrapping, following a procedure suggested by Stock and Watson (1996). Third, the performance of the Nyblom's statistic and of the F-tests could be affected by the presence of heteroskedasticity in the errors. Hence, we consider a robust version of these tests, where the relevant covariance matrices are estimated by the White's (1980) method.

From Table 5b, BIC selection does not affect the outcome of the tests. Instead, the use of finite sample critical values substantially decreases the fraction of unstable series according to Nyblom's statistic and the F-tests, values are now in the range 10-25%. The KP1, KP2 and BP tests are robust to the change of critical values, and still reject stability for about 20% of the series. The robust estimation of the covariance matrix further decreases the number of unstable series according to the F-tests, to values around 10%.

Overall, we can conclude that tests for parameter stability reject this hypothesis for about 20-30% of the series.

5.2 Forecast evaluation for unstable series

We now evaluate whether the time-varying and non-linear models under analysis forecast better than linear specifications for the unstable series. In particular, we select those series for which the Nyblom's test rejects at the 10% level, so that TVAR models should be particularly favored, see the Data Appendix for a list of these 134 variables.

Table 6 reports the ranking of the top-10 models, according to the loss function in equation (9). An AR in levels, with a constant, and either a fixed or AIC selected number of lags is still the best forecasting model on average. The main difference with respect to Table 2 is that now the TVAR3 ranks second when $h=1$ and $r=2$. More generally, TV models appear more frequently in the top-10, actually they are the only models in the top-10 not in the linear class, but still most of the top-10 models are linear. The ranking of STAR or ANN models does not improve significantly. This implies that, on average, linear models forecast comparatively well also in the presence of changing parameters.

Yet, if we increase the level of disaggregation by focusing on the percentage of series for which a model forecasts best in msfe terms, the outcome is rather different. From Table 7a, the improved performance of time-varying models for unstable series is evident, in particular for $h=3$ and 6. When $h=3$ the TV models forecast best for 51% of the unstable variables, versus 38% in Table 3a, and for 37% when $h=6$, versus 29% in Table 3a. The performance of the linear models deteriorates at all forecast horizons, in particular for $h=1,3$, while that of the non-linear models improves slightly for $h=1$ and 6 but deteriorates for $h=3$.

Within the TV class, the performance of the TV-AR improves for all forecast horizons, that for the STAR models for $h=3,6$. This pattern is also reflected in Table 7b, the fraction of series for which the TV models are among the top-N, with $N=5, 10, 15, 20$, increases for all values of N with respect to Table 3b.

We can further increase the level of disaggregation, by considering the performance of each model for each variable. In table 8 we report, for the best model in Table 6 and those in Table 7b, the mean of the empirical distribution of the msfe relative to the benchmark ARFC04, and some percentiles. If we focus on the mean, the ranking in Table 6 (for $r=2$) is obviously confirmed. But if we look at the percentiles, it is evident that the time-varying and non-linear models in Table 7b can perform substantially better than the benchmark and the ARFC0a for a consistent fraction of series. Moreover, the whole empirical distribution for the TV and non-linear models is in general shifted to the left with respect to the case where all the variables are analyzed, see Table 4, which is consistent with the fact that these models should perform better for unstable variables.

In summary, time-varying models forecast better for a larger fraction of unstable series, even though on average linear models are still the best.

6. Forecasting industrial production, unemployment and inflation

In this section we focus on three key macroeconomic variables, namely, industrial production (IP), unemployment (UNEM) and inflation (INFL), for all the 11 countries originally in the EMU. More precisely, we focus on IP growth, the change in the unemployment rate, and the growth in consumer prices.

In Table 9 we report the outcome of the stability tests for these variables. If we label unstable a variable when at least one test rejects, then inflation is the most unstable with 10 rejections out of 11, followed by unemployment with 5 out of 10 (unemployment is not available for Portugal over the whole sample), and IP ranks third with 5 rejections out of 11. Overall, there appears to be enough instability for the time-varying and non-linear models to be potentially useful.

In Table 10 we then compare the ranking of the competing models using the average loss function in equation (9) for several values of ρ , only the top-2 models are reported for each variable to save space. In the case of IP growth, the best models are linear, but for inflation an ANN model ranks second for $h=1$ and a TV model for $h=3$, while for unemployment time-varying models rank first (in particular, a TVAR for $h=1$ and STAR models for $h=3,6$). Hence, for this variable even the average msfe can be reduced by forecasting with a model outside the linear class.

Table 11 presents the ranking of the methods based on the fraction of series for which they msfe dominate. In this case time-varying methods perform rather well also for IP growth, in particular models in the STAR class and when $h=1,3$. The ANN models rank first in about 50% of the cases for unemployment, and they achieve good results also for inflation. Overall, linear models work well only in about 25% of the cases.

7. Conclusions

In this paper we have provided an accurate analysis of the relative forecasting performance of linear, time-varying and non-linear models, using a very large dataset of 480 macroeconomic variables for the countries in the European Monetary Union. The main finding is that models outside the linear class are useful for a substantial fraction of the series

Three final comments are in order to interpret this result. First, as shown for example in Clements and Hendry (1998, 1999), the best forecasting model does not necessarily provide the closest approximation to the generating mechanism of the process. Yet, the fact that time-varying and non-linear models work well for a substantial fraction of the series, combined with the outcome of the instability tests and with the

many social, economic and institutional changes that affected European countries over the past 20 years, provide enough evidence against the untested use of linear models.

Second, unfortunately, the sample available is rather short, so that estimation and forecast evaluation have to be conducted on a limited number of observations. We have compensated for this effect by increasing substantially the number of series under analysis. Moreover, a longer sample would have likely improved the estimation of non-linear and time-varying models, so that the results could point even more in their favor.

Third, other non-linear or time-varying models could perform even better, and computation of the forecasts using simulation methods could further increase the forecast gains.

Thus, this paper would like to encourage applied researchers to go beyond linearity, and economists to investigate in more detail the economic rationale underlying the good performance of non-linear models.

References

- Andrews, D.W.K. and Ploberger, W. (1994), "Optimal tests when a nuisance parameter is present only under the alternative", *Econometrica*, 62, 1383-1414.
- Artis, M. and Marcellino, M. (2001), "Fiscal forecasting: the track record of IMF, OECD", *Econometrics Journal*, 4, s20-s36.
- Box, G.E.P. and Jenkins, G.M. (1970), *Time series analysis, forecasting and control*, San Francisco: Holden Day.
- Breusch, T.S. and Pagan, A.R. (1979), "A simple test for heteroschedasticity and random coefficient variation", *Econometrica*, 47, 1287-1294.
- Clements, M.P. and Hendry, D.F. (1996), "Multi-step estimation for forecasting", *Oxford Bulletin of Economics and Statistics*, 58, 657-684.
- Clements, M.P. and Hendry, D.F. (1998), *Forecasting Economic Time Series*, Cambridge: Cambridge University Press.
- Clements, M.P. and Hendry, D.F. (1999), *Forecasting Non-Stationary Economic Time Series*, Cambridge (MA): MIT Press.
- Diebold, F.X. and Kilian, L. (2000), "Unit Root Tests are Useful for Selecting Forecasting Models," *Journal of Business and Economic Statistics*, 18, 265-273.
- Elliott, G., Rothenberg, T.J. and Stock, J.H. (1996), "Efficient tests for an autoregressive", *Econometrica*, 64, 813-36.
- Granger, C.W.J. and Newbold, P. (1986), *Forecasting economic time series*, San Diego: Academic Press.
- Granger, C.W.J. and Terasvirta, T. (1993), *Modelling nonlinear economic relationships*, Oxford: Oxford University Press.
- Hansen, B. (1992), "Tests for parameter instability in regressions with I(1) processes", *Journal of Business and Economic Statistics*, 10, 321-336.
- Hornik, K., Stinchcombe, M. and White, H. (1989), "Multilayer feedforward networks are universal approximators", *Neural Networks*, 2, 359-66.
- Marcellino, M., Stock, J.H. and Watson, M.W. (2000), "A dynamic factor analysis of the *mimeo*."
- Marcellino, M., Stock, J.H. and Watson, M.W. (2001), "Macroeconomic forecasting in the Euro area: country specific versus Euro wide information", *European Economic Review* (forthcoming).
- Makridakis, S. Anderson, A., Carbonne, R., Fildes, R., Hibon, M., Lewandowski, R., Newton, J., Parzen, E., Winkler, R. (1982), "The accuracy of extrapolation (time series) methods: Results of a forecasting competition", *Journal of Forecasting*, 1, 111-153.

- Meese, R. and Geweke, J. (1984), "A comparison of autoregressive univariate forecasting procedures for macroeconomic time series", *Journal of Business and Economic Statistics*, 2, 191-200.
- Nyblom, J. (1989), "Testing for constancy of parameters over time", *Journal of the American Statistical Association*, 84, 223-230.
- Ploberger, W. and Kramer, W. (1992), "The CUSUM test with OLS residuals", *Econometrica*, 60, 271-286.
- Quandt, R.E. (1960), "Tests of the hypothesis that a linear regression system obeys two", *Journal of the American Statistical Association*, 55, 324-330.
- Swanson, N.R. and White, H. (1997), "A model selection approach to real-time macroeconomic forecasting using linear models and artificial neural networks", *Review of Economics and Statistics*, 79, 540-550.
- Terasvirta, T. (1998), "Modelling economic relationships with smooth transition", Ullah, A. and Giles, D.E.A. (eds.), *Handbook of Applied Economic Statistics*, New York: Marcel Dekker, 507-552.
- Tiao, G.C. and Xu, D. (1993), "Robustness of maximum likelihood estimates for multi-step predictions: the exponential smoothing case", *Biometrika*, 80, 623-641.
- Stock, J.H. (1996), "VAR, error correction and pretest forecasts at long horizons", *Oxford Bulletin of Economics and Statistics*, 58, 685-701.
- Stock, J.H. and Watson, M.W. (1996), "Evidence on structural instability in macroeconomic time series relations", *Journal of Business and Economic Statistics*, 14, 11-30.
- Stock, J.H. and Watson, M.W. (1999), "A comparison of linear and nonlinear univariate models for forecasting macroeconomic time series", in Engle, R. and White, R. (eds), *Cointegration, causality, and forecasting: A festschrift in honor of Clive W.J. Granger*, Oxford: Oxford University Press, 1-44.
- White, H. (1980), "A heteroskedasticity consistent covariance matrix estimator and a direct test for heteroskedasticity", *Econometrica*, 48, 817-830.

Table 1 – Forecasting models

Linear methods

| | |
|------------|---|
| ARF(X,Y,Z) | <i>Autoregressive models</i> (18 models) X = C (const.) or T (trend) Y = 0 (stationary), 1 (I(1)), P (pre-test) Z = 4 (4 lags), a (AIC), b (BIC) |
| EX(X) | <i>Exponential smoothing</i> (3 models) X = 1 (single), 2 (double), P (pre-test) |
| NOCHANGE | <i>No change forecast</i> (1 model) |

Time-varying methods

| | |
|--------------|--|
| ARTVF(X,Y,Z) | <i>Time-varying AR models</i> (9 models) X = C (const.) Y = 0 (stationary), 1 (I(1)), P (pre-test) Z = 3 (3 lags), a (AIC), b (BIC) |
| LS(X,Y,Z) | <i>Logistic smooth transition</i> (6 models) X = 0 (stationary), 1 (I(1)), P (pre-test) Y = transition variable, 10 ($z_t = y_t$), 06 ($z_t = y_t - y_{t-6}$) Z = 3 (p, lag length) |
| LSF(X,W) | <i>Logistic smooth transition</i> (6 models) X = 0 (stationary), 1 (I(1)), P (pre-test) W = a (AIC on transition variable and p), b (BIC) |

Non-linear methods

| | |
|-------------|---|
| AN(X,Y,Z,W) | <i>Artificial neural network models</i> (9 models) X = 0 (stationary), 1 (I(1)), P (pre-test) Y = 2 (n_1) Z = 0, 1, 2 (n_2) W = 3 (p, lag length) |
| ANF(X,S) | <i>Artificial neural network models</i> (6 models) X = 0 (stationary), 1 (I(1)), P (pre-test) S = a (AIC on n_1, n_2, p), b (BIC) |

Table 2 - Ranking of competing models with different loss functions

| Rank | Horizon | $\rho=1$ | $\rho=1.5$ | $\rho=2$ | $\rho=2.5$ | $\rho=3$ |
|------|---------|----------|------------|----------|------------|----------|
| 1 | h=1 | ARFC04 | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| | h=3 | ARFC04 | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| | h=6 | ARFC0a | ARFC0a | ARFC0a | ARFC0a | ARFC0a |
| 2 | h=1 | ARFC0a | ARFC0a | ARFC0a | ARFC0a | ARFC0a |
| | h=3 | ARFC0a | ARFC0a | ARFC0a | ARFC0a | ARFC0a |
| | h=6 | ARFC04 | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| 3 | h=1 | ARFT1a | ARFT1a | ARFT1a | ARFC1a | ARFC1a |
| | h=3 | ARFC0b | ARFC0b | ARFC0b | ARFC0b | ARFC0b |
| | h=6 | ARFC0b | ARFC0b | ARFC0b | ARFC0b | ARFC0b |
| 4 | h=1 | ARFC0b | ARFC0b | ARFC1a | ARFTPa | ARFTPa |
| | h=3 | ARFC1a | ARFC1a | ARFC1a | ARFC1a | ARFC1a |
| | h=6 | ARFC1a | ARFC1a | ARTVFC03 | ARTVFC03 | ARTVFC03 |
| 5 | h=1 | ARFT1b | ARFC1a | ARFTPa | ARFT1a | ARFCPa |
| | h=3 | ARFTPa | ARFTPa | ARFTPa | ARFTPa | ARFTPa |
| | h=6 | ARFCPa | ARFCPa | ARFC1a | ARFT1a | ARTVFC0b |
| 6 | h=1 | ARFC1a | ARFTPa | ARFC0b | ARFC0b | ARFC0b |
| | h=3 | ARFCPa | ARFCPa | ARFCPa | ARFCPa | ARFCPa |
| | h=6 | ARFTPa | ARFTPa | ARFT1a | ARFC1a | ARFT1a |
| 7 | h=1 | ARFT14 | ARFT14 | ARFCPa | ARFCPa | ARFT1a |
| | h=3 | ARFC1b | ARFC1b | ARTVFC1b | ARTVFC03 | ARTVFC03 |
| | h=6 | ARFT1a | ARFT1a | ARFCPa | ARFCPa | ARTVFC0a |
| 8 | h=1 | ARFTPa | ARFT1b | ARFT14 | ARTVFC1b | ARTVFC1b |
| | h=3 | ARFTPb | ARTVFC1b | ARTVFCPb | ARTVFC1b | ARTVFC1b |
| | h=6 | ARFC1b | ARTVFC03 | ARFTPa | ARFTPa | ARFC1a |
| 9 | h=1 | ARTVFC03 | ARFCPa | ARTVFC1b | ARTVFC1a | ARTVFC1a |
| | h=3 | ARFT1a | ARFT1a | ARFC1b | ARTVFCPb | ARTVFCPb |
| | h=6 | ARTVFC1a | ARFT14 | ARFT14 | ARTVFC0b | ARFCPa |
| 10 | h=1 | ARFCPa | ARTVFC1b | ARTVFC1a | ARFT14 | ARTVFCP3 |
| | h=3 | ARTVFC1b | ARFTPb | ARTVFC03 | ARTVFC1a | ARTVFCPa |
| | h=6 | ARFCPb | ARFC1b | ARFT1b | ARFT1b | ARFT1b |

Notes:

See Table1 for definition of models

The loss function is $Loss_m^h = \frac{1}{N} \sum_{n=1}^N \frac{Loss_{n,m}^h}{Loss_{n,1}^h}$, $Loss_{n,m}^h = \frac{1}{T-h} \sum_{t=1}^{T-h} |e_{t+h,n,m}|^r$, where the

benchmark model is ARFC04 and e_{t+h} is the h-step ahead forecast error

Table 3a – Fraction of series for which a forecasting method has lowest msfe

| Method | AR | ES | NoChange | ARTV | LSTAR | ANN |
|--------|------|------|----------|------|-------|------|
| h=1 | 0.23 | 0.11 | 0.02 | 0.13 | 0.25 | 0.27 |
| h=3 | 0.20 | 0.09 | 0.02 | 0.10 | 0.28 | 0.27 |
| h=6 | 0.23 | 0.11 | 0.03 | 0.06 | 0.23 | 0.32 |

| Method | Linear | Time-varying | Non-linear |
|--------|--------|--------------|------------|
| h=1 | 0.36 | 0.38 | 0.27 |
| h=3 | 0.31 | 0.38 | 0.27 |
| h=6 | 0.37 | 0.29 | 0.32 |

Notes:

Linear method includes AR,ES, No Change.

Time-varying method includes ARTV, LSTAR.

Non-linear method includes ANN.

Figures do not sum up to one because of rounding errors.

Table 3b – Fraction of series for which a forecasting model is in the top-N

| | N=1 | N=5 | N=10 | N=15 | N=20 |
|----------|------|------|------|------|------|
| ARFT0b | 0.03 | 0.12 | 0.22 | 0.32 | 0.39 |
| | 0.03 | 0.14 | 0.21 | 0.31 | 0.37 |
| | 0.04 | 0.14 | 0.19 | 0.26 | 0.32 |
| EX1 | 0.05 | 0.16 | 0.25 | 0.30 | 0.36 |
| | 0.05 | 0.15 | 0.24 | 0.30 | 0.35 |
| | 0.05 | 0.16 | 0.25 | 0.29 | 0.36 |
| NOCHANGE | 0.02 | 0.08 | 0.12 | 0.16 | 0.21 |
| | 0.02 | 0.09 | 0.15 | 0.19 | 0.24 |
| | 0.03 | 0.13 | 0.19 | 0.23 | 0.27 |
| ARTVFC03 | 0.04 | 0.14 | 0.24 | 0.35 | 0.44 |
| | 0.03 | 0.16 | 0.28 | 0.37 | 0.41 |
| | 0.02 | 0.07 | 0.21 | 0.31 | 0.38 |
| LS0063 | 0.04 | 0.12 | 0.18 | 0.26 | 0.34 |
| | 0.04 | 0.13 | 0.22 | 0.30 | 0.37 |
| | 0.04 | 0.11 | 0.21 | 0.30 | 0.37 |
| ANF0b | 0.05 | 0.13 | 0.20 | 0.26 | 0.30 |
| | 0.04 | 0.14 | 0.22 | 0.28 | 0.33 |
| | 0.05 | 0.14 | 0.22 | 0.27 | 0.32 |

Notes: See Table 1 for definition of models

The figures report the fraction of series for which a model is among the N models with the lowest msfe.

The reported models are the best performers in each class for N=1.

For each model, the three rows report result for, respectively, h=1,3,6.

Table 4 – Mean and percentiles of relative msfe for selected forecasting models

| Forecast | Mean | 0.02 | 0.10 | 0.25 | 0.50 | 0.75 | 0.90 | 0.98 |
|----------|------|------|------|------|------|------|------|-------|
| ARFC04 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| ARFC0a | 1.01 | 0.85 | 0.94 | 0.99 | 1.00 | 1.03 | 1.08 | 1.19 |
| | 1.01 | 0.77 | 0.92 | 0.98 | 1.00 | 1.04 | 1.08 | 1.21 |
| | 1.00 | 0.66 | 0.84 | 0.94 | 1.00 | 1.03 | 1.14 | 1.38 |
| ARFT0b | 1.08 | 0.67 | 0.87 | 0.95 | 1.02 | 1.11 | 1.30 | 1.93 |
| | 1.22 | 0.45 | 0.74 | 0.88 | 1.04 | 1.26 | 1.79 | 3.44 |
| | 1.84 | 0.19 | 0.49 | 0.75 | 1.08 | 1.84 | 2.99 | 9.79 |
| EX1 | 1.50 | 0.65 | 0.85 | 0.94 | 1.05 | 1.26 | 2.02 | 5.00 |
| | 1.81 | 0.43 | 0.76 | 0.91 | 1.06 | 1.40 | 3.08 | 9.75 |
| | 3.55 | 0.16 | 0.44 | 0.76 | 1.07 | 1.78 | 5.24 | 20.32 |
| NOCHANGE | 1.67 | 0.64 | 0.88 | 1.02 | 1.24 | 1.56 | 2.24 | 5.00 |
| | 1.90 | 0.42 | 0.77 | 0.95 | 1.22 | 1.63 | 3.04 | 9.75 |
| | 3.65 | 0.15 | 0.44 | 0.81 | 1.27 | 2.18 | 4.56 | 18.28 |
| ARTVFC03 | 1.03 | 0.72 | 0.90 | 0.96 | 1.01 | 1.07 | 1.17 | 1.52 |
| | 1.07 | 0.59 | 0.81 | 0.92 | 1.01 | 1.12 | 1.34 | 2.14 |
| | 1.24 | 0.37 | 0.62 | 0.86 | 1.04 | 1.29 | 1.96 | 3.94 |
| LS0063 | 1.14 | 0.68 | 0.88 | 0.97 | 1.05 | 1.16 | 1.36 | 2.04 |
| | 1.15 | 0.55 | 0.81 | 0.95 | 1.05 | 1.21 | 1.60 | 2.52 |
| | 1.36 | 0.27 | 0.60 | 0.84 | 1.03 | 1.35 | 2.23 | 4.93 |
| ANF0b | 1.31 | 0.66 | 0.88 | 0.98 | 1.08 | 1.24 | 1.53 | 2.15 |
| | 1.42 | 0.41 | 0.78 | 0.95 | 1.14 | 1.47 | 2.09 | 4.76 |
| | 2.75 | 0.16 | 0.48 | 0.87 | 1.30 | 2.15 | 4.13 | 12.17 |

Notes:

The models are the best in Table 2 and those in Table 3b

The benchmark model is ARFC04

For each forecast the three rows correspond to, respectively, $h=1, 3, 6$

See Table 1 for the definition of the models

Table 5a –Stability Tests

Percentage of series significant at

| | NY | KP1 | KP2 | BP | QLR | MLR | ALR |
|-----------|------|------|------|------|------|------|------|
| 10% level | 27.9 | 22.5 | 20.6 | 31.5 | 40.8 | 34.6 | 41.7 |
| 5% level | 21.0 | 16.0 | 16.9 | 26.7 | 35.2 | 28.5 | 35.2 |
| 1% level | 8.3 | 5.2 | 9.6 | 17.3 | 24.6 | 17.5 | 23.3 |

Notes:

The model is an AR3, except for the second line where the lag length is selected by BIC

NY is Nyblom's (1989) test

KP1 and KP2 are the Ploebeger and Kramer's (1992) supremum and mean square tests

BP is Breusch and Pagan's (1979) Lagrange multiplier test

QLR is Quandt's (1960) supremum F-test

MLR is Andrews and Ploebeger's (1994) mean F-test

ALR is Andrews and Ploebeger's (1994) average exponential F-test

Table 5b –Stability Tests, sensitivity analysis

Percentage of series significant at the 10%

| Lag L | Cval | H-R | NY | KP1 | KP2 | BP | QLR | MLR | ALR |
|-------|------|-----|------|------|------|------|------|------|------|
| Fixed | Asy. | No | 27.9 | 22.5 | 20.6 | 31.5 | 40.8 | 34.6 | 41.7 |
| BIC | Asy. | No | 25.4 | 21.9 | 20.4 | 33.5 | 41.7 | 33.8 | 39.6 |
| Fixed | MC | No | 10.6 | 24.8 | 21.3 | 20.0 | 25.6 | 16.5 | 23.8 |
| Fixed | MC | Yes | 8.8 | -- | -- | -- | 10.8 | 10.6 | 11.3 |

Notes:

Lag L indicates whether the lag length is fixed at 3 or BIC determined

Cval indicates whether the critical values are asymptotic or for finite sample (MC)

H-R indicates robust estimation of the relevant covariance matrices

See the Notes to Table 5a for a description of the tests

Table 6 - Unstable series , ranking of competing models with different loss functions

| Rank | | $\rho=1$ | $\rho=1.5$ | $\rho=2$ | $\rho=2.5$ | $\rho=3$ |
|------|-----|----------|------------|----------|------------|----------|
| 1 | h=1 | ARFT1b | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| | h=3 | ARFT1a | ARFC0a | ARFC0a | ARFC0a | ARFC04 |
| | h=6 | ARFC0a | ARFC0a | ARFC0a | ARFC0a | ARFC0a |
| 2 | h=1 | ARFT1a | ARTVFC03 | ARTVFC03 | ARFC0a | ARFC0a |
| | h=3 | ARFC0a | ARFT1a | ARFC04 | ARFC04 | ARFC0a |
| | h=6 | ARFC04 | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| 3 | h=1 | ARFC04 | ARFT1a | ARFT1a | ARFC0b | ARFC0b |
| | h=3 | ARFT1b | ARFC04 | ARFC0b | ARFC0b | ARFC0b |
| | h=6 | ARFC0b | ARFC0b | ARFC0b | ARFC0b | ARFC0b |
| 4 | h=1 | ARTVFC03 | ARFT1b | ARFC0a | ARFT14 | ARFT14 |
| | h=3 | ARFT14 | ARFC0b | ARFT1a | ARFT1a | ARFT1a |
| | h=6 | ARFT1a | ARFT1a | ARFT1a | ARFT1a | ARFT1a |
| 5 | h=1 | ARFT14 | ARFT14 | ARFT14 | ARFT1a | ARFT1a |
| | h=3 | EXP | ARFT1b | ARFT14 | ARTVFC03 | ARTVFC03 |
| | h=6 | ARFT14 | ARFT14 | ARFT14 | ARFT14 | ARFT14 |
| 6 | h=1 | ARFC0b | ARFC0a | ARFT1b | ARTVFC03 | ARFT1b |
| | h=3 | ARFC0b | ARFT14 | ARFT1b | ARFT14 | ARFT14 |
| | h=6 | ARFT1b | ARFT1b | ARFT1b | ARFT1b | ARFT1b |
| 7 | h=1 | ARFC0a | ARFC0b | ARFC0b | ARFT1b | ARTVFC03 |
| | h=3 | ARFC04 | ARTVFC03 | ARTVFC03 | ARFT1b | ARFT1b |
| | h=6 | ARTVFC03 | ARTVFC03 | ARTVFC03 | ARTVFC03 | ARTVFC03 |
| 8 | h=1 | EXP | EXP | EXP | ARTVFC1b | ARTVFC1b |
| | h=3 | ARTVFC03 | EXP | EXP | ARTVFC0b | ARTVFC0b |
| | h=6 | EXP | ARTVFC0b | ARTVFC0b | ARTVFC0b | ARTVFC0b |
| 9 | h=1 | EX2 | EX2 | ARTVFC1b | ARTVFC1a | ARTVFCpb |
| | h=3 | EX2 | ARTVFC0b | ARTVFC0b | ARTVFC0a | ARTVFC0a |
| | h=6 | EX2 | ARTVFC0a | ARTVFC0a | ARTVFC0a | ARTVFC0a |
| 10 | h=1 | ARTVFC0a | ARTVFC1b | ARTVFC1a | ARTVFCpb | ARTVFC1a |
| | h=3 | ARTVFC0b | ARTVFC0a | ARTVFC0a | EXP | EXP |
| | h=6 | ARFC1a | EXP | EXP | EXP | EXP |

Notes:

See Table1 for definition of models

The loss function is $Loss^h_m = \frac{1}{N} \sum_{n=1}^N \frac{Loss^h_{n,m}}{Loss^h_{n,1}}$, $Loss^h_{n,m} = \frac{1}{T-h} \sum_{t=1}^{T-h} |e_{t+h,n,m}|^r$, where the

benchmark model is ARFC04 and e_{t+h} is the h-step ahead forecast error

Table 7a – Fraction of unstable series for which a forecasting method has lowest msfe

| Method | AR | ES | NoChange | ARTV | LSTAR | ANN |
|--------|------|------|----------|------|-------|------|
| h=1 | 0.17 | 0.04 | 0.04 | 0.17 | 0.21 | 0.31 |
| h=3 | 0.15 | 0.1 | 0.01 | 0.16 | 0.35 | 0.20 |
| h=6 | 0.22 | 0.08 | 0.05 | 0.08 | 0.29 | 0.28 |

| Method | Linear | Time-varying | Non-linear |
|--------|--------|--------------|------------|
| h=1 | 0.25 | 0.38 | 0.31 |
| h=3 | 0.17 | 0.51 | 0.20 |
| h=6 | 0.35 | 0.37 | 0.28 |

Notes:

Linear method includes AR,ES, No Change.

Time-varying method includes ARTV, LSTAR.

Non-linear method includes ANN.

Figures do not sum up to one because of rounding errors.

Table 7b – Fraction of unstable series for which a forecasting model is in the top-N

| | N=1 | N=5 | N=10 | N=15 | N=20 |
|----------|------|------|------|------|------|
| ARFT1b | 0.02 | 0.12 | 0.30 | 0.39 | 0.54 |
| | 0.04 | 0.14 | 0.28 | 0.35 | 0.48 |
| | 0.04 | 0.17 | 0.25 | 0.34 | 0.43 |
| EX1 | 0.01 | 0.13 | 0.20 | 0.25 | 0.34 |
| | 0.06 | 0.13 | 0.22 | 0.25 | 0.30 |
| | 0.04 | 0.14 | 0.21 | 0.24 | 0.31 |
| NOCHANGE | 0.04 | 0.13 | 0.18 | 0.24 | 0.29 |
| | 0.01 | 0.11 | 0.16 | 0.20 | 0.26 |
| | 0.05 | 0.13 | 0.21 | 0.24 | 0.28 |
| ARTVFC03 | 0.09 | 0.23 | 0.34 | 0.49 | 0.57 |
| | 0.06 | 0.19 | 0.33 | 0.46 | 0.51 |
| | 0.03 | 0.07 | 0.23 | 0.34 | 0.40 |
| LS0103 | 0.03 | 0.13 | 0.25 | 0.34 | 0.40 |
| | 0.07 | 0.25 | 0.33 | 0.39 | 0.46 |
| | 0.07 | 0.21 | 0.31 | 0.36 | 0.41 |
| ANF0b | 0.09 | 0.21 | 0.29 | 0.37 | 0.39 |
| | 0.04 | 0.18 | 0.27 | 0.34 | 0.40 |
| | 0.05 | 0.18 | 0.29 | 0.36 | 0.41 |

Notes: See Table 1 for definition of models

The figures report the fraction of series for which a model is among the N models with the lowest msfe.

The reported models are the best performers in each class for N=1.

Table 8 – Unstable series, mean and percentiles of relative msfe for selected forecasting models

| Forecast | Mean | 0.02 | 0.10 | 0.25 | 0.50 | 0.75 | 0.90 | 0.98 |
|----------|------|------|------|------|------|------|------|-------|
| ARFC04 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 | 1.00 |
| ARFC0a | 1.01 | 0.87 | 0.97 | 0.99 | 1.00 | 1.03 | 1.07 | 1.20 |
| | 0.99 | 0.74 | 0.89 | 0.97 | 1.00 | 1.04 | 1.08 | 1.20 |
| | 0.98 | 0.62 | 0.81 | 0.91 | 0.98 | 1.03 | 1.15 | 1.37 |
| ARFT1b | 1.01 | 0.69 | 0.84 | 0.95 | 1.01 | 1.06 | 1.18 | 1.45 |
| | 1.03 | 0.34 | 0.62 | 0.87 | 1.03 | 1.17 | 1.34 | 2.06 |
| | 1.29 | 0.14 | 0.35 | 0.68 | 1.01 | 1.29 | 2.74 | 5.19 |
| EX1 | 1.59 | 0.70 | 0.86 | 0.95 | 1.09 | 1.48 | 2.14 | 4.57 |
| | 1.91 | 0.39 | 0.74 | 0.92 | 1.15 | 1.86 | 3.84 | 7.42 |
| | 2.83 | 0.15 | 0.40 | 0.79 | 1.36 | 2.63 | 6.23 | 18.26 |
| NOCHANGE | 1.67 | 0.64 | 0.85 | 1.00 | 1.14 | 1.63 | 2.36 | 4.57 |
| | 1.94 | 0.36 | 0.73 | 0.94 | 1.18 | 1.96 | 3.84 | 7.42 |
| | 2.86 | 0.13 | 0.39 | 0.79 | 1.38 | 2.64 | 6.23 | 18.26 |
| ARTVFC03 | 1.01 | 0.68 | 0.85 | 0.92 | 0.99 | 1.05 | 1.17 | 1.53 |
| | 1.03 | 0.56 | 0.75 | 0.87 | 0.97 | 1.10 | 1.29 | 2.57 |
| | 1.33 | 0.28 | 0.50 | 0.83 | 1.03 | 1.47 | 2.05 | 3.94 |
| LS0103 | 1.15 | 0.68 | 0.88 | 0.96 | 1.06 | 1.18 | 1.45 | 1.89 |
| | 1.19 | 0.46 | 0.73 | 0.87 | 1.05 | 1.35 | 1.73 | 2.89 |
| | 1.67 | 0.12 | 0.40 | 0.70 | 1.09 | 1.68 | 3.46 | 8.21 |
| ANF0b | 1.13 | 0.67 | 0.82 | 0.94 | 1.05 | 1.21 | 1.41 | 2.15 |
| | 1.40 | 0.33 | 0.73 | 0.90 | 1.13 | 1.41 | 2.07 | 4.54 |
| | 2.57 | 0.08 | 0.33 | 0.66 | 1.15 | 2.63 | 6.28 | 14.97 |

Notes:

The benchmark model is ARFC04

The models are the best from Table 6 and those from Table 7b

For each forecast the three rows correspond to, respectively, h=1, 3, 6

See Table 1 for the definition of the models

Table 9 –Stability Tests for Unemployment, IP and Inflation

| Series | NY | KP1 | KP2 | BP | QLR | MLR | ALR | | |
|----------|------|----------|---------|----------|-----------|-----------|-----------|--|-----|
| Ger-UNEM | 0.86 | 1.08 | 0.25 | 5.98 | 11.76 | 5.38 | 3.33 | | |
| Ger-IP | 0.78 | 1.14 | 0.19 | 4.85 | 12.11 | 4.63 | 3.87 | | |
| Ger-INFL | 1.03 | 1.13 | 0.3 | 8.48 | ** 13.6 | 7.92 | ** 4.89 | | * |
| Ita-UNEM | 0.47 | 0.87 | 0.17 | 10.29 | ** 8.71 | 3.78 | 2.41 | | |
| Ita-IP | 0.57 | 0.58 | 0.07 | 2.15 | 8.37 | 3.3 | 2.13 | | |
| Ita-INFL | 0.78 | 0.73 | 0.21 | 7.54 | * 11.23 | 5.3 | 3.08 | | |
| Spa-UNEM | 0.83 | 0.79 | 0.16 | 1.34 | 19.65 | ** 6.89 | * 6.81 | | ** |
| Spa-IP | 1.34 | ** 0.99 | 0.14 | 5.76 | 16.21 | ** 8.21 | ** 5.53 | | ** |
| Spa-INFL | 1.74 | *** 1.51 | ** 1.07 | *** 9.52 | ** 39.79 | *** 15.01 | *** 15.76 | | *** |
| Fra-UNEM | 0.34 | 0.66 | 0.08 | 3.44 | 6.89 | 2.06 | 1.5 | | |
| Fra-IP | 0.52 | 0.64 | 0.09 | 7.43 | * 6.69 | 2.95 | 1.82 | | |
| Fra-INFL | 1.3 | ** 1.04 | 0.44 | * 3.18 | 30.39 | *** 15.45 | *** 11.48 | | *** |
| Aus-UNEM | 0.4 | 0.69 | 0.06 | 9.16 | ** 9 | 2.4 | 1.91 | | |
| Aus-IP | 0.8 | 0.64 | 0.07 | 0.78 | 8.85 | 4.59 | 2.79 | | |
| Aus-INFL | 0.39 | 0.94 | 0.26 | 0.78 | 9.34 | 2.71 | 2.33 | | |
| Lux-UNEM | 0.92 | 1.23 | * 0.38 | * 7.59 | * 9.96 | 5.33 | 3.22 | | |
| Lux-IP | 0.58 | 0.98 | 0.24 | 5.1 | 9.52 | 3.23 | 2.38 | | |
| Lux-INFL | 0.77 | 1.08 | 0.19 | 5.32 | 19.47 | ** 4.89 | 5.7 | | ** |
| Net-UNEM | 1.99 | *** 0.75 | 0.13 | 9.26 | ** 23.88 | *** 11.03 | *** 7.4 | | *** |
| Net-IP | 0.49 | 0.48 | 0.02 | 5.54 | 7.16 | 3.49 | 2.17 | | |
| Net-INFL | 1.28 | ** 1.26 | * 0.34 | * 4.17 | 20.85 | *** 10.5 | *** 7.33 | | *** |
| Fin-UNEM | 1.92 | *** 0.85 | 0.14 | 8.65 | ** 135.35 | *** 54.7 | *** 63.69 | | *** |
| Fin-IP | 1.27 | ** 1.31 | * 0.31 | 3.45 | 17.09 | ** 8.71 | ** 5.66 | | ** |
| Fin-INFL | 1.41 | ** 1.5 | ** 0.78 | *** 3.06 | 20.2 | *** 12.47 | *** 7.6 | | *** |
| Por-IP | 0.46 | 1.29 | * 0.23 | 1.59 | 8.79 | 2.66 | 2.15 | | |
| Por-INFL | 1.23 | ** 1.15 | 0.57 | ** 27.29 | *** 38.04 | *** 11.05 | *** 14.85 | | *** |
| Bel-UNEM | 0.66 | 0.98 | 0.18 | 4.65 | 8.93 | 4.47 | 2.76 | | |
| Bel-IP | 0.32 | 0.72 | 0.08 | 11.21 | ** 5.44 | 1.48 | 0.89 | | |
| Bel-INFL | 1.66 | *** 1.48 | ** 0.5 | ** 12.35 | *** 35.86 | *** 14.38 | *** 13.41 | | *** |
| Ire-UNEM | 0.63 | 0.83 | 0.23 | 1.58 | 11.06 | 4.55 | 3.02 | | |
| Ire-IP | 0.69 | 0.98 | 0.28 | 1.9 | 7.77 | 4.13 | 2.34 | | |
| Ire-INFL | 1.4 | ** 0.93 | 0.2 | 0.98 | 16.13 | ** 8.49 | ** 6.14 | | ** |

Notes:

The model is an AR3 for the change in unemployment, and growth in IP and CPI

NY is Nyblom's (1989) test

KP1 and KP2 are the Ploberger and Kramer's (1992) supremum and mean square tests

BP is Breusch and Pagan's (1979) Lagrange multiplier test

QLR is Quandt's (1960) supremum F-test

MLR is Andrews and Ploberger's (1994) mean F-test

ALR is Andrews and Ploberger's (1994) average exponential F-test

*, **, and *** indicate significance at, respectively, 10%, 5% and 1% level

Table 10 - Ranking of competing models with different loss functions, selected series

IP growth

| Rank | | $\rho=1$ | $\rho=1.5$ | $\rho=2$ | $\rho=2.5$ | $\rho=3$ |
|------|-----|----------|------------|----------|------------|----------|
| 1 | h=1 | ARFTP4 | ARFTP4 | ARFTP4 | ARFTP4 | ARFCP4 |
| | h=3 | ARFTP4 | ARFCP4 | ARFTP4 | ARFCP4 | ARFCP4 |
| | h=6 | ARFCPa | ARFCP4 | ARFTP4 | ARFTP4 | ARTVFC13 |
| 2 | h=1 | ARFC14 | ARFC14 | ARFC14 | ARFC14 | ARFC14 |
| | h=3 | ARFC14 | ARFC14 | ARFC14 | ARFC14 | ARFC14 |
| | h=6 | ARFTPa | ARFC14 | ARFC14 | ARFC14 | ARTVFCP3 |

Unemployment (change)

| Rank | | $\rho=1$ | $\rho=1.5$ | $\rho=2$ | $\rho=2.5$ | $\rho=3$ |
|------|-----|----------|------------|----------|------------|----------|
| 1 | h=1 | NOCHANGE | ARTVFCP3 | ARTVFCP3 | ARTVFC13 | ARTVFC13 |
| | h=3 | LSP063 | LS1063 | LS1063 | LSP063 | LSP063 |
| | h=6 | ANP213 | ANP213 | LSF1a | ANF1b | ANFPb |
| 2 | h=1 | ARFTPb | ARTVFC13 | ARTVFC13 | ARTVFCP3 | ARTVFCP3 |
| | h=3 | LS1063 | LSP063 | LSP063 | LS1063 | LS1063 |
| | h=6 | AN1213 | AN1213 | LSFPa | ANFPb | ANF1b |

CPI inflation

| Rank | | $\rho=1$ | $\rho=1.5$ | $\rho=2$ | $\rho=2.5$ | $\rho=3$ |
|------|-----|----------|------------|----------|------------|----------|
| 1 | h=1 | ARFC04 | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| | h=3 | LS0103 | ARFC04 | ARFC04 | ARFC04 | ARFC04 |
| | h=6 | AN0223 | ARFC04 | ARFC04 | ARFC0a | ARFC0a |
| 2 | h=1 | ANF0b | ANF0b | ANF0b | ANF0b | ANF0b |
| | h=3 | ARFC04 | LS0103 | LS0103 | LS0103 | LS0103 |
| | h=6 | ARFC04 | AN0223 | ARFC0a | ARFC04 | ARFC04 |

Notes:

See Table1 for definition of models

The loss function is $Loss^h_m = \frac{1}{N} \sum_{n=1}^N \frac{Loss^h_{n,m}}{Loss^h_{n,1}}$, $Loss^h_{n,m} = \frac{1}{T-h} \sum_{t=1}^{T-h} |e_{t+h,n,m}|^r$, where the

benchmark model is ARFC04 and e_{t+h} is the h-step ahead forecast error

Unemployment for Portugal is not available

Table 11 - Fraction of series for which a forecasting method has lowest msfe

IP growth

| Method | AR | ES | NoChange | ARTV | LSTAR | ANN |
|--------|------|------|----------|------|-------|------|
| h=1 | 2/11 | - | - | 1/11 | 7/11 | 1/11 |
| h=3 | 1/11 | 1/11 | - | 2/11 | 6/11 | 1/11 |
| h=6 | 4/11 | - | - | 1/11 | - | 6/11 |

| Method | Linear | Time-varying | Non-linear |
|--------|--------|--------------|------------|
| h=1 | 2/11 | 8/11 | 1/11 |
| h=3 | 2/11 | 8/11 | 1/11 |
| h=6 | 4/11 | 1/11 | 6/11 |

Unemployment (change)

| Method | AR | ES | NoChange | ARTV | LSTAR | ANN |
|--------|------|------|----------|------|-------|------|
| h=1 | 3/10 | 1/10 | - | 2/10 | - | 4/10 |
| h=3 | 2/10 | - | - | - | 3/10 | 5/10 |
| h=6 | 1/10 | 1/10 | - | - | 3/10 | 5/10 |

| Method | Linear | Time-varying | Non-linear |
|--------|--------|--------------|------------|
| h=1 | 4/10 | 2/10 | 4/10 |
| h=3 | 2/10 | 3/10 | 5/10 |
| h=6 | 2/10 | 3/10 | 5/10 |

CPI inflation

| Method | AR | ES | NoChange | ARTV | LSTAR | ANN |
|--------|------|------|----------|------|-------|------|
| h=1 | 3/11 | 1/11 | - | 1/11 | - | 6/11 |
| h=3 | 2/11 | - | - | 2/11 | 5/11 | 2/11 |
| h=6 | 1/11 | - | 1/11 | 1/11 | 3/11 | 5/11 |

| Method | Linear | Time-varying | Non-linear |
|--------|--------|--------------|------------|
| h=1 | 4/11 | 1/11 | 6/11 |
| h=3 | 2/11 | 7/11 | 2/11 |
| h=6 | 2/11 | 4/11 | 5/11 |

Notes:

Unemployment for Portugal is not available

Linear method includes AR,ES, No Change.

Time-varying method includes ARTV, LSTAR.

Non-linear method includes ANN.

Appendix: The dataset

The first column reports the OECD identifier of the series. The second column reports ***, **, * when the Nyblom (1989) test for parameter stability rejects at the, respectively, 1%, 5%, and 10% level. The third column reports a brief description of the series.

Austria

| OECD Code | OECD Definition |
|-----------|--|
| 7020349K | * Consumer goods, sa /Industrial production /PRODUCTION 1990=100 Austria /AUTNSO-OECD STATI |
| 7020439K | Intermediate goods, sa /Industrial production /PRODUCTION 1990=100 Austria /AUTNSO-OECD ST |
| 7020449K | Investment goods, sa /Industrial production /PRODUCTION 1990=100 Austria /AUTNSO-OECD STAT |
| 7020519K | Total, sa /Industrial production /PRODUCTION 1990=100 Austria /AUTNSO-OECD STATISTICS, PAF |
| 70206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Austria /INTISI-OECD STATISTICS, PAR |
| 7032419K | Total: value, sa /Retail sales /DOMESTIC TRADE 1990=100 Austria /AUTNSO-OECD STATISTICS, F |
| 7032439K | Durable goods: value, sa /Retail sales /DOMESTIC TRADE 1990=100 Austria /AUTNSO-OECD STAT |
| 7032449K | RETAIL SALES (volume), sa 1990 = 100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7032519K | Total: value, sa /Wholesale sales /DOMESTIC TRADE 1990=100 Austria /AUTNSO-OECD STATISTIC |
| 70325383 | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Austria / |
| 70426780 | Foreign workers /Employment /LABOUR '000 Austria /AUTLAB-OECD STATISTICS, PARIS" |
| 70428283 | Registered unemployed, sa /Unemployment /LABOUR '000 Austria /AUTLAB-OECD STATISTICS, PA |
| 704284A3 | Rate, sa /Unemployment /LABOUR % Austria /AUTLAB-OECD STATISTICS, PARIS" |
| 70429983 | *** Unfilled vacancies, sa /Labour - other /LABOUR (continued) '000 Austria /AUTLAB-OECD STATI |
| 7043119H | Hourly rates /Wages /WAGES 1990=100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7043219K | Monthly earnings, sa /Wages /WAGES 1990=100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7043779H | PRODUCER PRICES (manufacturing) 1990 = 100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7044029H | ** Agricultural goods /Wholesale prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PAF |
| 7044119H | Food /Wholesale prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7044219H | * Petroleum products /Wholesale prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PA |
| 7044259H | *** Transport equipment /Wholesale prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, P, |
| 7044459H | Food /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7044479H | Fuel and electricity /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PA |
| 7044559H | All items less food /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PAF |
| 7044579H | All items less food less rent /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STATI |
| 7044589H | Rent /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7044619H | All items /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STATISTICS, PARIS" |
| 7044639H | All items excl. seasonal items /Consumer prices /PRICES 1990=100 Austria /AUTNSO-OECD STAT |
| 7054821D | AUT MONETARY AGGREGATE M1 SA /MN SCHILLING Austria OECD STATISTICS, PARIS" |
| 7054829D | MONETARY AGGREGATES, sa 1990 = 100 Austria /AUTCBA-OECD STATISTICS, PARIS" |
| 7054831D | ** AUT MONETARY AGGREGATE (M3) SA /MN SCHILLING Austria OECD STATISTICS, PARIS" |
| 7054839D | ** MONETARY AGGREGATES, sa 1990 = 100 Austria /AUTCBA-OECD STATISTICS, PARIS" |
| 7054911A | * AUT SAVINGS DEPOSITS /MN SCHILLING Austria OECD STATISTICS, PARIS" |
| 7054911X | AUT FOREIGN EXCHANGE DEPOSITS /MN SCHILLING Austria OECD STATISTICS, PARIS" |
| 7055111A | ** AUT QUASI-MONEY /MN SCHILLING Austria OECD STATISTICS, PARIS" |
| 7055251A | ** Domestic credit /Domestic finance /DOMESTIC FINANCE S bln Austria /AUTCBA-OECD STATISTICS |
| 705561AH | Official discount /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Austria /AUTCBA-OE |
| 7055809H | VSE WBI Index /Share prices /INTEREST RATES - SHARE PRICES 1990=100 Austria /AUTSTE-OEC |
| 705581AH | Yield of public sector bonds /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Austria |
| 7056009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Austria /OECD-OECD STATISTICS, PARIS" |
| 705601AH | EXCHANGE RATES National currency units per US dollar Austria /OECD-OECD STATISTICS, PARIS |
| 705611AS | ** Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Austria /INTIMF |
| 7056151A | Net foreign position /Foreign finance /FOREIGN FINANCE S bln Austria /AUTCBA-OECD STATISTI |
| 70663200 | Current account balance /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCBA-C |
| 70663250 | AUT BOP CURRENT BALANCE /MN US DOLLARS Austria OECD STATISTICS, PARIS" |
| 70663400 | Net current transfers /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCBA-OEC |
| 70663500 | Financial account balance /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCBA |
| 70663600 | Net services /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCBA-OECD STATI |
| 70663700 | Net errors and omissions /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCBA- |
| 70663900 | ** Change in official reserves /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCB |
| 70664000 | Net investment income /Balance of payments /BALANCE OF PAYMENTS S mln Austria /AUTCBA-OE |
| 70765103 | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE S bln Austria /AUTNSO-OECD STA |
| 70765253 | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Aust |
| 70765303 | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE S bln Austria /AUTNSO-OECD STATISTICS, PA |
| 70765553 | FOREIGN TRADE - Ftr Imports (fob/cif) Total, sa Billions US dollars; monthly averages Aust |
| 70765603 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE S bln Austria /AUTNSO-OECD STATISTICS, PA |
| 70765753 | FOREIGN TRADE - Ftr Exports Fob Total, sa Billions US dollars; monthly averages Austria /A |

Belgium

| OECD Code | OECD Definition |
|-----------|--|
| 2220339K | Construction, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO-OECD STATISTI |
| 2220359K | Consumer durable goods, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO-OEC |
| 2220369K | Consumer non-durable goods, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO |
| 2220439K | Intermediate goods, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO-OECD ST |
| 2220449K | Investment goods, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO-OECD STAT |
| 2220459K | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO-OECD STATIS |
| 2220519K | Total, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELNSO-OECD STATISTICS, PAR |
| 2220539K | Total including construction, sa /Industrial production /PRODUCTION 1990=100 Belgium /BELN |
| 22206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Belgium /INTISI-OECD STATISTICS, PAR |
| 2232048X | BEL CON BUILDING STARTED RESID /CUB METERS Belgium OECD STATISTICS, PARIS" |
| 22321180 | Total /Permits issued /CONSTRUCTION cu. m. '000 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 22321283 | Residential, sa /Permits issued /CONSTRUCTION cu. m. '000 Belgium /BELNSO-OECD STATISTICS, |
| 22321480 | Total /Buildings started /CONSTRUCTION cu. m. '000 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 22321780 | CONSTRUCTION Thousands; monthly averages Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2232419K | Total: value, sa /Retail sales /DOMESTIC TRADE 1990=100 Belgium /BELNSO-OECD STATISTICS, P |
| 2232449Y | Total: volume, sa /Retail sales /DOMESTIC TRADE 1990=100 Belgium /BELNSO-OECD STATISTICS, |
| 22325383 | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Belgium / |
| 224280A3 | Rate, sa /Unemployment /LABOUR % Belgium /BELLAB-OECD STATISTICS, PARIS" |
| 22428183 | Total, sa /Unemployment /LABOUR '000 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 224284A0 | * BEL UNEMPLOY. % CIV. LAB. FORCE /PERCNT Belgium OECD STATISTICS, PARIS" |
| 224284AX | BEL UNEMPL % INSURED LAB FORCE /PERCNT Belgium OECD STATISTICS, PARIS" |
| 224286A3 | STANDARDISED UNEMPLOYMENT RATES, sa Per cent Belgium /INTEUR-OECD STATISTICS, PARIS" |
| 2243459H | *** Chemicals /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2243749H | ** Consumer goods /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2243529H | Food, beverages and tobacco /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATIS |
| 2243649H | Intermediate goods /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARI |
| 2243659H | ** Investment goods /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2243749H | Petroleum products /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARI |
| 2243759H | Textiles and clothing /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, P |
| 2243779H | Manufactured goods /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARI |
| 2243869H | Total /Producer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2244449H | BEL CPI ENERGY /I/90 Belgium OECD STATISTICS, PARIS" |
| 2244459H | *** Food /Consumer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2244479H | Fuel and electricity /Consumer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PA |
| 2244499H | All goods less food /Consumer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PAR |
| 2244559H | * BEL CPI NON FOOD /I/90 Belgium OECD STATISTICS, PARIS" |
| 2244589H | ** Rent /Consumer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2244599H | *** Services less rent /Consumer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARI |
| 2244619H | *** All items /Consumer prices /PRICES 1990=100 Belgium /BELNSO-OECD STATISTICS, PARIS" |
| 2254829D | MONETARY AGGREGATES, sa 1990 = 100 Belgium /BELCBA-OECD STATISTICS, PARIS" |
| 2254839D | ** MONETARY AGGREGATES, sa 1990 = 100 Belgium /BELCBA-OECD STATISTICS, PARIS" |
| 225567AH | Treasury certificates /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Belgium /BELCB |
| 225578AH | Yield of government bonds /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Belgium /B |
| 2256009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Belgium /OECD-OECD STATISTICS, PARIS" |
| 225601AH | EXCHANGE RATES National currency units per US dollar Belgium /OECD-OECD STATISTICS, PARIS" |
| 225611AS | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Belgium /INTIMF |
| 22765103 | ** Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE FB bln Belgium /BELNSO-OECD ST |
| 22765303 | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE FB bln Belgium /BELNSO-OECD STATISTICS, P |
| 22765603 | *** Exports f.o.b., sa /Foreign trade /FOREIGN TRADE FB bln Belgium /BELNSO-OECD STATISTICS, P |

Finland

| OECD Code | OECD Definition |
|-----------|---|
| 6420349J | Consumer goods, sa /Industrial production /PRODUCTION 1990=100 Finland /FINNSO-OECD STATIS |
| 6420439J | Intermediate goods, sa /Industrial production /PRODUCTION 1990=100 Finland /FINNSO-OECD ST |
| 6420449J | * Investment goods, sa /Industrial production /PRODUCTION 1990=100 Finland /FINNSO-OECD STAT |
| 6420459J | *** Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Finland /FINNSO-OECD STATIS |
| 6420519J | ** Total, sa /Industrial production /PRODUCTION 1990=100 Finland /FINNSO-OECD STATISTICS, PAR |
| 64206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Finland /INTISI-OECD STATISTICS, PAR |
| 64207182 | Wood fellings, sa /Commodity output /PRODUCTION cu. m. mln Finland /FINNSO-OECD STATISTICS |
| 64321180 | Total /Permits issued /CONSTRUCTION cu. m. mln Finland /FINNSO-OECD STATISTICS, PARIS" |

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| 64321280 | Residential /Permits issued /CONSTRUCTION cu. m. mln Finland /FINNSO-OECD STATISTICS, PARI |
| 6432239H | ** Total /Cost of construction /CONSTRUCTION 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6432449J | ** Volume, sa /Retail sales /DOMESTIC TRADE 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6432519J | Value, sa /Wholesale sales /DOMESTIC TRADE 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS |
| 64325383 | * New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Finland / |
| 6432589J | Volume, sa /Wholesale sales /DOMESTIC TRADE 1990=100 Finland /FINNSO-OECD STATISTICS, PARI |
| 64426580 | ** FIN EMPLOYMENT TOTAL /PERSONS Finland OECD STATISTICS, PARIS" |
| 6442659H | ** TOTAL EMPLOYMENT 1990 = 100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 64426883 | ** FIN EMPLOYMENT INDUSTRY SA /PERSONS Finland OECD STATISTICS, PARIS" |
| 64427480 | * Part-time (economic reasons) /Employment /LABOUR '000 Finland /FINNSO-OECD STATISTICS, PAR |
| 644280A2 | Rate, sa /Unemployment /LABOUR % Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6442819J | *** Total, sa /Unemployment /LABOUR '000 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 64429180 | *** Total hours worked: industry /Labour - other /LABOUR hrs mln Finland /FINNSO-OECD STATISTI |
| 64429983 | Unfilled vacancies, sa /Labour - other /LABOUR '000 Finland /FINNSO-OECD STATISTICS, PARIS |
| 6443479H | ** Consumer goods /Producer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6443649H | Intermediate goods /Producer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARI |
| 6443659H | Investment goods /Producer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6443749H | ** Petroleum products /Producer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARI |
| 6443869H | PRODUCER PRICES (manufacturing) 1990 = 100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6444419H | ** Beverages and tobacco /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, P |
| 6444459H | ** Food /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6444479H | ** Fuel and electricity /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PA |
| 6444509H | ** All items less food less housing /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD ST |
| 6444529H | *** Housing /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6444559H | ** All items less food /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PAR |
| 6444619H | ** All items /Consumer prices /PRICES 1990=100 Finland /FINNSO-OECD STATISTICS, PARIS" |
| 6444709H | ** FIN CPI NON FOOD NON ENERGY /I/90 Finland OECD STATISTICS, PARIS" |
| 6454821D | Monetary aggregate (M1), sa /Domestic finance /DOMESTIC FINANCE Fmk bln Finland /FINCBA-OE |
| 6454829D | MONETARY AGGREGATES, sa 1990 = 100 Finland /FINCBA-OECD STATISTICS, PARIS" |
| 6454831D | *** Monetary aggregate (M3), sa /Domestic finance /DOMESTIC FINANCE Fmk bln Finland /FINCBA-OE |
| 6454839D | *** MONETARY AGGREGATES, sa 1990 = 100 Finland /FINCBA-OECD STATISTICS, PARIS" |
| 6454841B | *** Monetary aggregate (M2), sa /Domestic finance /DOMESTIC FINANCE Fmk bln Finland /FINCBA-OE |
| 6455231A | Credit to economy /Domestic finance /DOMESTIC FINANCE Fmk bln Finland /FINCBA-OECD STATIST |
| 645561AH | Base rate /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Finland /FINCBA-OECD STATI |
| 6455631H | Liquidity credit rate /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Finland /FINCB |
| 6455849H | HEX All Share Index /Share prices /INTEREST RATES - SHARE PRICES 1990=100 Finland /FINCBA- |
| 6456009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Finland /OECD-OECD STATISTICS, PARIS" |
| 645601AH | EXCHANGE RATES National currency units per US dollar Finland /OECD-OECD STATISTICS, PARIS" |
| 645611AS | * Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Finland /INTIMF |
| 64663100 | *** Trade balance /Balance of payments /BALANCE OF PAYMENTS Fmk bln Finland /FINCBA-OECD STATI |
| 64663200 | * Current account balance /Balance of payments /BALANCE OF PAYMENTS Fmk bln Finland /FINCBA- |
| 64663400 | Net current transfers /Balance of payments /BALANCE OF PAYMENTS Fmk bln Finland /FINCBA-OE |
| 64663500 | Financial account balance /Balance of payments /BALANCE OF PAYMENTS Fmk bln Finland /FINCB |
| 64663700 | Net errors and omissions /Balance of payments /BALANCE OF PAYMENTS Fmk bln Finland /FINCBA |
| 64664000 | Net investment income /Balance of payments /BALANCE OF PAYMENTS Fmk bln Finland /FINCBA-OE |
| 64765303 | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE Fmk bln Finland /FINNSO-OECD STATISTICS, |
| 64765603 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE Fmk bln Finland /FINNSO-OECD STATISTICS, |

France

| OECD Code | OECD Definition |
|-----------|---|
| 1420339J | Construction, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STATISTIC |
| 1420349J | Consumer goods, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STATIST |
| 1420399J | Energy, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STATISTICS, PAR |
| 1420439J | Intermediate goods, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STA |
| 1420449J | Investment goods, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STATI |
| 1420459J | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STATISTI |
| 1420519J | Total, sa /Industrial production /PRODUCTION 1990=100 France /FRANSO-OECD STATISTICS, PARI |
| 14206183 | Passenger cars, sa /Commodity output /PRODUCTION '000 France /FRAND-OECD STATISTICS, PARI |
| 14206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 France /INTISI-OECD STATISTICS, PARI |
| 14321780 | CONSTRUCTION Thousands; monthly averages France /FRATRA-OECD STATISTICS, PARIS" |
| 1432419J | *** Value, sa /Retail sales /DOMESTIC TRADE 1990=100 France /FRACHA-OECD STATISTICS, PARIS" |
| 1432449J | Volume, sa /Retail sales /DOMESTIC TRADE 1990=100 France /FRACHA-OECD STATISTICS, PARIS" |
| 14325382 | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 France /F |
| 1432549J | Manufact. products - 1980 prices, sa /Retail sales /DOMESTIC TRADE 1990=100 France /FRANSO |
| 14428282 | Registered unemployed, sa /Unemployment /LABOUR '000 France /FRALAB-OECD STATISTICS, PARIS |
| 144286A3 | STANDARDISED UNEMPLOYMENT RATES, sa Per cent France /INTEUR-OECD STATISTICS, PARIS" |

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|----------|---|
| 1442879J | New jobseekers, sa /Unemployment /LABOUR 1990=100 France /FRALAB-OECD STATISTICS, PARIS" |
| 1443249H | *** Labour cost: engineering industries /Wages /WAGES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1443259H | *** Labour cost: textile industries /Wages /WAGES 1990=100 France /FRANSO-OECD STATISTICS, PAR |
| 1443419J | Agricultural goods, sa /Producer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, P |
| 1443459H | Chemicals /Producer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1443649H | Intermediate goods /Producer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS |
| 1443699H | Metal products /Producer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1443809H | FRA WPI INTERM PRICE OF RAW MATER //90 France OECD STATISTICS, PARIS" |
| 1444449H | FRA CPI ENERGY //90 France OECD STATISTICS, PARIS" |
| 1444459H | *** Food /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1444479H | Fuel and electricity /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PAR |
| 1444499H | *** All goods less food /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARI |
| 1444559H | ** All items less food /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARI |
| 1444589H | *** Rent /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1444599H | *** Services less rent /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS |
| 1444619H | ** All items /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1444659H | ** Paris: all items /Consumer prices /PRICES 1990=100 France /FRANSO-OECD STATISTICS, PARIS" |
| 1454822B | ** Monetary aggregate (M1), sa /Domestic finance /DOMESTIC FINANCE FF bln France /FRACBA-OECD |
| 1454829B | ** MONETARY AGGREGATES, sa 1990 = 100 France /FRACBA-OECD STATISTICS, PARIS" |
| 1454832B | ** Monetary aggregate (M3), sa /Domestic finance /DOMESTIC FINANCE FF bln France /FRACBA-OECD |
| 1454839B | ** MONETARY AGGREGATES, sa 1990 = 100 France /FRACBA-OECD STATISTICS, PARIS" |
| 1454892B | Investment aggregate (P1), sa /Domestic finance /DOMESTIC FINANCE FF bln France /FRACBA-OE |
| 1455631H | Call money /Interest rates /INTEREST RATES - SHARE PRICES % p.a. France /FRACBA-OECD STATI |
| 145565AH | 3-month PIBOR /Interest rates /INTEREST RATES - SHARE PRICES % p.a. France /FRACBA-OECD ST |
| 145581AH | Bonds: public and semi-public /Interest rates /INTEREST RATES - SHARE PRICES % p.a. France |
| 1455849H | Paris Stock Exchange: SBF 250 /Share prices /INTEREST RATES - SHARE PRICES 1990=100 France |
| 1456009H | EFFECTIVE EXCHANGE RATES 1990 = 100 France /OECD-OECD STATISTICS, PARIS" |
| 145601AH | EXCHANGE RATES National currency units per US dollar France /OECD-OECD STATISTICS, PARIS" |
| 145611AS | ** Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln France /INTIMF- |
| 14765102 | Net trade (f.o.b.-f.o.b.), sa /Foreign trade /FOREIGN TRADE FF bln France /FRACUS-OECD STA |
| 14765252 | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Fran |
| 14765302 | Imports f.o.b., sa /Foreign trade /FOREIGN TRADE FF bln France /FRACUS-OECD STATISTICS, PA |
| 14765552 | FOREIGN TRADE - Ftr Imports (fob/cif) Total, sa Billions US dollars; monthly averages Fran |
| 14765602 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE FF bln France /FRACUS-OECD STATISTICS, PA |
| 14765752 | FOREIGN TRADE - Ftr Exports Fob Total, sa Billions US dollars; monthly averages France /FR |

Germany

| OECD Code | OECD Definition |
|-----------|--|
| 1220519J | INDUSTRIAL PRODUCTION, sa 1990 = 100 Germany /DEUCBA-OECD STATISTICS, PARIS" |
| 12206180 | Passenger cars /Commodity output /PRODUCTION '000 Germany /DEUNSO-OECD STATISTICS, PARIS" |
| 12206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Germany /INTISI-OECD STATISTICS, PAR |
| 12321100 | Total /Permits issued /CONSTRUCTION DM bln Germany /DEUNSO-OECD STATISTICS, PARIS" |
| 12321200 | Residential /Permits issued /CONSTRUCTION DM bln Germany /DEUNSO-OECD STATISTICS, PARIS" |
| 1232449K | ** RETAIL SALES (volume), sa 1990 = 100 Germany /OECD-OECD STATISTICS, PARIS" |
| 12325383 | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Germany / |
| 1242669K | DEU CIVILIAN EMPLOYMENT SA //90 Germany OECD STATISTICS, PARIS" |
| 12427183 | Manufacturing, sa /Employment /LABOUR '000 Germany /DEUNSO-OECD STATISTICS, PARIS" |
| 12427480 | Part-time (economic reasons) /Employment /LABOUR '000 Germany /DEUNSO-OECD STATISTICS, PAR |
| 12428280 | Registered unemployed /Unemployment /LABOUR '000 Germany /DEULAB-OECD STATISTICS, PARIS" |
| 124286A3 | STANDARDISED UNEMPLOYMENT RATES, sa -- ADJUSTED Down by 2% in xxx (AC) Per cent Germany |
| 12429180 | Monthly hours of work /Labour - other /LABOUR hrs mln Germany /DEULAB-OECD STATISTICS, PAR |
| 12430082 | Unfilled vacancies, sa /Labour - other /LABOUR '000 Germany /DEUCBA-OECD STATISTICS, PARIS |
| 1243569H | PRODUCER PRICES (manufacturing) 1990 = 100 Germany /DEUNSO-OECD STATISTICS, PARIS" |
| 1244619H | CONSUMER PRICES 1990 = 100 Germany /DEUNSO-OECD STATISTICS, PARIS" |
| 1254821B | Monetary aggregate (M1), sa /Domestic finance /DOMESTIC FINANCE DM bln Germany /DEUCBA-OEC |
| 1254829B | MONETARY AGGREGATES, sa 1990 = 100 Germany /DEUCBA-OECD STATISTICS, PARIS" |
| 1254829D | DEU MONETARY AGGT M1 RFA+RDA EST SA //90 Germany OECD STATISTICS, PARIS" |
| 1254831B | Monetary aggregate (M3), sa /Domestic finance /DOMESTIC FINANCE DM bln Germany /DEUCBA-OEC |
| 1254839B | MONETARY AGGREGATES, sa 1990 = 100 Germany /DEUCBA-OECD STATISTICS, PARIS" |
| 1254839D | DEU M1 + QUASI MONEY RFA+RDA(EST)SA //90 Germany OECD STATISTICS, PARIS" |
| 1254841B | Monetary aggregate (M2), sa /Domestic finance /DOMESTIC FINANCE DM bln Germany /DEUCBA-OEC |
| 1254911A | Personal savings deposits /Domestic finance /DOMESTIC FINANCE DM bln Germany /DEUCBA-OECD |
| 1254931B | Monetary aggregate (M3+), sa /Domestic finance /DOMESTIC FINANCE DM bln Germany /DEUCBA-OE |
| 1255231D | Credit to economy, sa /Domestic finance /DOMESTIC FINANCE DM bln Germany /DEUCBA-OECD STAT |
| 125561AH | Official discount /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Germany /DEUCBA-OE |
| 1255631H | * Call money /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Germany /DEUCBA-OECD STAT |
| 125565AH | 3-month FIBOR /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Germany /DEUCBA-OECD S |

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| 125581AH | Public sector bond yield /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Germany /DE |
| 1255849H | CDAX Share Price Index /Share prices /INTEREST RATES - SHARE PRICES 1990=100 Germany /DEUN |
| 1256009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Germany /OECD-OECD STATISTICS, PARIS" |
| 125611AS | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Germany /INTIMF |
| 1256151A | Net foreign position /Foreign finance /FOREIGN FINANCE DM bln Germany /DEUCBA-OECD STATIST |
| 12663200 | Current account balance /Balance of payments /BALANCE OF PAYMENTS DM bln Germany /DEUCBA-O |
| 12663250 | FDR/DEU BOP CURRENT BALANCE /MN US DOLLARS Germany OECD STATISTICS, PARIS" |
| 12663500 | Financial account balance /Balance of payments /BALANCE OF PAYMENTS DM bln Germany /DEUCBA |
| 12663700 | Net errors and omissions /Balance of payments /BALANCE OF PAYMENTS DM bln Germany /DEUCBA- |
| 12663900 | Change in official reserves /Balance of payments /BALANCE OF PAYMENTS DM bln Germany /DEUC |
| 12765102 | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE DM bln Germany /DEUNSO-OECD ST |
| 12765252 | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Germ |
| 12765302 | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE DM bln Germany /DEUNSO-OECD STATISTICS, P |
| 12765552 | FOREIGN TRADE - Ftr Imports (fob/cif) Total, sa Billions US dollars; monthly averages Germ |
| 12765602 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE DM bln Germany /DEUNSO-OECD STATISTICS, P |
| 12765752 | FOREIGN TRADE - Ftr Exports Fob Total, sa Billions US dollars; monthly averages Germany /D |

Ireland

| OECD Code | OECD Definition |
|-----------|---|
| 2820349J | Consumer goods, sa /Industrial production /PRODUCTION 1990=100 Ireland /IRLNSO-OECD STATIS |
| 2820439J | Intermediate goods, sa /Industrial production /PRODUCTION 1990=100 Ireland /IRLNSO-OECD ST |
| 2820449J | Investment goods, sa /Industrial production /PRODUCTION 1990=100 Ireland /IRLNSO-OECD STAT |
| 2820459J | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Ireland /IRLNSO-OECD STATIST |
| 2820519J | Total, sa /Industrial production /PRODUCTION 1990=100 Ireland /IRLNSO-OECD STATISTICS, PAR |
| 2832249H | ** Residential /Cost of construction /CONSTRUCTION 1990=100 Ireland /IRLENV-OECD STATISTICS, |
| 2832419J | Value, sa /Retail sales /DOMESTIC TRADE 1990=100 Ireland /IRLNSO-OECD STATISTICS, PARIS" |
| 2832449J | ** Volume, sa /Retail sales /DOMESTIC TRADE 1990=100 Ireland /IRLNSO-OECD STATISTICS, PARIS" |
| 28325383 | ** New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Ireland / |
| 28427480 | Part-time (economic reasons) /Employment /LABOUR '000 Ireland /IRLNSO-OECD STATISTICS, PAR |
| 28428282 | Registered unemployed, sa /Unemployment /LABOUR '000 Ireland /IRLNSO-OECD STATISTICS, PARI |
| 284286A3 | STANDARDISED UNEMPLOYMENT RATES, sa Per cent Ireland /INTEUR-OECD STATISTICS, PARIS" |
| 2844049H | Investment goods /Wholesale prices /PRICES 1990=100 Ireland /IRLNSO-OECD STATISTICS, PARIS |
| 2844119H | Food /Wholesale prices /PRICES 1990=100 Ireland /IRLNSO-OECD STATISTICS, PARIS" |
| 2844189H | Manufactured goods /Wholesale prices /PRICES 1990=100 Ireland /IRLNSO-OECD STATISTICS, PAR |
| 2844269H | ** Total /Wholesale prices /PRICES 1990=100 Ireland /IRLNSO-OECD STATISTICS, PARIS" |
| 2854829D | *** MONETARY AGGREGATES, sa 1990 = 100 Ireland /IRLCBA-OECD STATISTICS, PARIS" |
| 2855631H | Call money /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Ireland /IRLCBA-OECD STAT |
| 2855809H | ISEQ Index - Overall /Share prices /INTEREST RATES - SHARE PRICES 1990=100 Ireland /IRLCBA |
| 2856009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Ireland /OECD-OECD STATISTICS, PARIS" |
| 285601AH | EXCHANGE RATES National currency units per US dollar Ireland /OECD-OECD STATISTICS, PARIS" |
| 285611AS | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Ireland /INTIMF |
| 28765102 | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE pdlr mln Ireland /IRLNSO-OECD |
| 28765252 | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Ire |
| 28765302 | * Imports c.i.f., sa /Foreign trade /FOREIGN TRADE pdlr mln Ireland /IRLNSO-OECD STATISTICS, |
| 28765552 | FOREIGN TRADE - Ftr Imports (fob/cif) Total, sa Billions US dollars; monthly averages Ire |
| 28765602 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE pdlr mln Ireland /IRLNSO-OECD STATISTICS, |
| 28765752 | FOREIGN TRADE - Ftr Exports Fob Total, sa Billions US dollars; monthly averages Ireland /I |

Italy

| OECD Code | OECD Definition |
|-----------|--|
| 1620349J | Consumer goods, sa /Industrial production /PRODUCTION 1990=100 Italy /ITANSO-OECD STATISTI |
| 1620439J | Industrial materials, sa /Industrial production /PRODUCTION 1990=100 Italy /ITANSO-OECD ST |
| 1620449J | Investment goods, sa /Industrial production /PRODUCTION 1990=100 Italy /ITANSO-OECD STATIS |
| 1620459J | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Italy /ITANSO-OECD STATIS |
| 1620519J | Total, sa /Industrial production /PRODUCTION 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS |
| 16206180 | Passenger cars /Commodity output /PRODUCTION '000 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 16206480 | Commercial vehicles /Commodity output /PRODUCTION '000 Italy /ITANSO-OECD STATISTICS, PARI |
| 16206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Italy /INTISI-OECD STATISTICS, PARIS |
| 1631299H | * Consumer goods /Sales /MANUFACTURING 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1631309H | Intermediate goods /Sales /MANUFACTURING 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1631319H | Investment goods /Sales /MANUFACTURING 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1631329H | Total /Sales /MANUFACTURING 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |

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| 1632019H | * | Total /New orders /MANUFACTURING 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1632249H | *** | Residential /Cost of construction /CONSTRUCTION 1990=100 Italy /ITANSO-OECD STATISTICS, PA |
| 1632419K | * | Major outlets: value, sa /Retail sales /DOMESTIC TRADE 1990=100 Italy /ITANSO-OECD STATIST |
| 1632449K | | RETAIL SALES (volume), sa 1990 = 100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 16325383 | | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Italy /IT |
| 164286A3 | | STANDARDISED UNEMPLOYMENT RATES, sa Per cent Italy /INTEUR-OECD STATISTICS, PARIS" |
| 16429880 | * | Labour disputes: time lost /Labour - other /LABOUR hrs '000 Italy /ITANSO-OECD STATISTICS, |
| 1643119H | ** | Hourly rates /Wages /WAGES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1643429H | * | Machinery and equipment /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, P |
| 1643459H | *** | Chemical products /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1643529H | | Food,, beverages and tobacco /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTI |
| 1643709H | *** | Non-metallic mineral products /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATIST |
| 1643719H | | Metal and metal products /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, |
| 1643749H | | Petroleum products /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1643759H | | Textiles and clothing /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PAR |
| 1643869H | | Total /Producer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1644419H | | Beverages and tobacco /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PAR |
| 1644459H | * | Food /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1644489H | | Fuel and electricity /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARI |
| 1644499H | ** | All goods less food /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS |
| 1644559H | | All items less food /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS |
| 1644589H | | Rent /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1644599H | | Services less rent /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1644619H | | All items /Consumer prices /PRICES 1990=100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1644679H | | CONSUMER PRICES 1990 = 100 Italy /ITANSO-OECD STATISTICS, PARIS" |
| 1654822D | *** | Monetary aggregate (M1), sa /Domestic finance /DOMESTIC FINANCE Lit '000 bln Italy /ITACBA |
| 1654832A | ** | ITA TOTAL LIQUIDITY /BN ITA LIRA Italy OECD STATISTICS, PARIS" |
| 1654833D | ** | Monetary aggregate (M2), sa /Domestic finance /DOMESTIC FINANCE Lit '000 bln Italy /ITACBA |
| 1654839D | ** | MONETARY AGGREGATES, sa 1990 = 100 Italy /ITACBA-OECD STATISTICS, PARIS" |
| 165498AH | ** | 3-month interbank deposits /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Italy /IT |
| 1655121A | | Gross bond issues: public sector /Domestic finance /DOMESTIC FINANCE Lit '000 bln Italy /I |
| 1655131A | | Gross bond issues: banking sector /Domestic finance /DOMESTIC FINANCE Lit '000 bln Italy / |
| 1655251A | * | Domestic credit /Domestic finance /DOMESTIC FINANCE Lit '000 bln Italy /ITACBA-OECD STATIS |
| 1655292A | | Finance to the non state sector /Domestic finance /DOMESTIC FINANCE Lit '000 bln Italy /IT |
| 1655751H | | Bond yield /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Italy /ITACBA-OECD STATIS |
| 165578AH | | Long-term treasury bonds /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Italy /ITAC |
| 1655849H | | ISE MIB Storico /Share prices /INTEREST RATES - SHARE PRICES 1990=100 Italy /ITACBA-OECD S |
| 1656009H | | EFFECTIVE EXCHANGE RATES 1990 = 100 Italy /OECD-OECD STATISTICS, PARIS" |
| 165601AH | | EXCHANGE RATES National currency units per US dollar Italy /OECD-OECD STATISTICS, PARIS" |
| 165611AS | | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Italy /INTIMF-O |
| 1656152A | | Net foreign position /Foreign finance /FOREIGN FINANCE Lit '000 bln Italy /ITACBA-OECD STA |
| 16663100 | | Trade balance /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /ITACBA-OECD ST |
| 16663200 | | Current account balance /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /ITAC |
| 16663250 | | ITA BOP CURRENT BALANCE US \$ /MN US \$ Italy OECD STATISTICS, PARIS" |
| 16663400 | | Net current transfers /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /ITACBA |
| 16663500 | | Financial account balance /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /IT |
| 16663600 | | Net services /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /ITACBA-OECD STA |
| 16663700 | | Net errors and omissions /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /ITA |
| 16663900 | | Change in official reserves /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy / |
| 16664000 | | Net income /Balance of payments /BALANCE OF PAYMENTS Lit '000 bln Italy /ITACBA-OECD STATI |
| 16765103 | | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE Lit bln Italy /ITASCO-OECD STA |
| 16765253 | | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Ital |
| 16765303 | * | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE Lit bln Italy /ITASCO-OECD STATISTICS, PA |
| 16765553 | ** | FOREIGN TRADE - Ftr Imports (fob/cif) Total, sa Billions US dollars; monthly averages Ital |
| 16765603 | | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE Lit bln Italy /ITASCO-OECD STATISTICS, PA |
| 16765753 | * | FOREIGN TRADE - Ftr Exports Fob Total, sa Billions US dollars; monthly averages Italy /ITA |

Luxembourg

| OECD Code | OECD Definition |
|-----------|--|
| 2420339K | Construction, sa /Industrial production /PRODUCTION 1990=100 Luxembourg /LUXNSO-OECD STATI |
| 2420459K | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Luxembourg /LUXNSO-OECD STAT |
| 2420519K | Total, sa /Industrial production /PRODUCTION 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, |
| 24206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Luxembourg /INTISI-OECD STATISTICS, |
| 24321383 | Permits issued, sa /Construction /CONSTRUCTION number Luxembourg /LUXNSO-OECD STATISTICS, |
| 24325383 | New passenger car registrations, sa /Domestic trade /DOMESTIC TRADE number Luxembourg /LUX |
| 2442699H | Industry: employees /Employment /LABOUR 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, PARIS |

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|----------|--|
| 24427080 | Iron and steel: wage earners /Employment /LABOUR '000 Luxembourg /LUXNSO-OECD STATISTICS, |
| 24428283 | Registered unemployed, sa /Unemployment /LABOUR number Luxembourg /LUXNSO-OECD STATISTICS, |
| 244286A3 | STANDARDISED UNEMPLOYMENT RATES, sa Per cent Luxembourg /INTEUR-OECD STATISTICS, PAR |
| 2442929H | Monthly hours of work /Labour - other /LABOUR 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, |
| 24429983 | Unfilled vacancies, sa /Labour - other /LABOUR number Luxembourg /OECD-OECD STATISTICS, PA |
| 2443159H | Monthly earnings /Wages /WAGES 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, PARIS" |
| 2443589H | Industrial goods /Producer prices /PRICES 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, PAR |
| 2444459H | Food /Consumer prices /PRICES 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, PARIS" |
| 2444479H | Fuel and electricity /Consumer prices /PRICES 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, |
| 2444559H | All items less food /Consumer prices /PRICES 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, |
| 2444619H | All items /Consumer prices /PRICES 1990=100 Luxembourg /LUXNSO-OECD STATISTICS, PARIS" |

Netherlands

| OECD Code | OECD Definition |
|-----------|--|
| 1820459J | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Netherlands /NLDNSO-OECD STA |
| 1820519J | Total, sa /Industrial production /PRODUCTION 1990=100 Netherlands /NLDNSO-OECD STATISTICS, |
| 18206680 | ** Crude petroleum /Commodity output /PRODUCTION tonnes '000 Netherlands /NLDNSO-OECD STATIST |
| 18206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Netherlands /INTISI-OECD STATISTICS, |
| 18206880 | Natural gas /Commodity output /PRODUCTION cu. m. mln Netherlands /NLDNSO-OECD STATISTICS, |
| 18321100 | Total /Permits issued /CONSTRUCTION f. mln Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 18321203 | Residential, sa /Permits issued /CONSTRUCTION f. mln Netherlands /NLDNSO-OECD STATISTICS, |
| 1832419K | Total: value, sa /Retail sales /DOMESTIC TRADE 1990=100 Netherlands /NLDNSO-OECD STATISTIC |
| 1832449K | RETAIL SALES (volume), sa 1990 = 100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 18325383 | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Netherlan |
| 184286A3 | *** STANDARDISED UNEMPLOYMENT RATES, sa Per cent Netherlands /INTEUR-OECD STATISTICS, PARIS" |
| 1843149H | Hourly rates: manufacturing /Wages /WAGES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, PA |
| 1843469H | ** Output: consumer goods /Producer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTI |
| 1843489H | Output: crude petroleum /Producer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATIST |
| 1843569H | ** PRODUCER PRICES (manufacturing) 1990 = 100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1843649H | * Output: intermediate goods /Producer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STAT |
| 1843659H | Output: investment goods /Producer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATI |
| 1843879H | ** Input: total /Producer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1843889H | * Output: total /Producer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, PARIS |
| 1844459H | ** Food /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1844479H | Fuel and electricity /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS |
| 1844499H | All goods less food /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, |
| 1844559H | All items less food /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, |
| 1844589H | Rent /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1844599H | ** Services less rent /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, |
| 1844619H | ** All items /Consumer prices /PRICES 1990=100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1844709H | NLD CPI NON FOOD-NON ENERGY //90 Netherlands OECD STATISTICS, PARIS" |
| 1854829D | MONETARY AGGREGATES, sa 1990 = 100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1854832D | NLD MONETARY AGGREGATE M3 SA /MN GUILDER Netherlands OECD STATISTICS, PARIS" |
| 1854839D | MONETARY AGGREGATES, sa 1990 = 100 Netherlands /NLDNSO-OECD STATISTICS, PARIS" |
| 1855631H | Call money (Amsterdam) /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Netherlands / |
| 1856009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Netherlands /OECD-OECD STATISTICS, PARIS" |
| 185601AH | EXCHANGE RATES National currency units per US dollar Netherlands /OECD-OECD STATISTICS, PA |
| 185611AS | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Netherlands /IN |
| 1856151A | *** Net foreign position /Foreign finance /FOREIGN FINANCE f. mln Netherlands /NLDNSO-OECD STA |
| 18765103 | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE f. mln Netherlands /NLDNSO-OEC |
| 18765253 | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Neth |
| 18765303 | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE f. mln Netherlands /NLDNSO-OECD STATISTIC |
| 18765603 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE f. mln Netherlands /NLDNSO-OECD STATISTIC |

Portugal

| OECD Code | OECD Definition |
|-----------|---|
| 3620459K | Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Portugal /PRTNSO-OECD STATIS |
| 3620519K | Total, sa /Industrial production /PRODUCTION 1990=100 Portugal /PRTNSO-OECD STATISTICS, PA |
| 36206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Portugal /INTISI-OECD STATISTICS, PA |
| 36428280 | Registered unemployed /Unemployment /LABOUR '000 Portugal /PRTEPT-OECD STATISTICS, PARIS" |
| 36429980 | ** Unfilled vacancies /Labour - other /LABOUR '000 Portugal /PRTEPT-OECD STATISTICS, PARIS" |
| 3644459H | ** Food /Consumer prices /PRICES 1990=100 Portugal /PRTNSO-OECD STATISTICS, PARIS" |

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| 3644549H | *** | Lisbon: all items less rent /Consumer prices /PRICES 1990=100 Portugal /PRTNSO-OECD STATIS |
| 3644559H | *** | All items less food and rent /Consumer prices /PRICES 1990=100 Portugal /PRTNSO-OECD STATI |
| 3644609H | ** | All items less rent /Consumer prices /PRICES 1990=100 Portugal /PRTNSO-OECD STATISTICS, PA |
| 3654829D | | MONETARY AGGREGATES, sa 1990 = 100 Portugal /PRTCBA-OECD STATISTICS, PARIS" |
| 3654831D | | PRT MONETARY AGGREGATE M2- SA /MN ESCUDO Portugal OECD STATISTICS, PARIS" |
| 3654839D | | MONETARY AGGREGATES, sa 1990 = 100 Portugal /PRTCBA-OECD STATISTICS, PARIS" |
| 3654861A | * | Total liquidity (L-) /Domestic finance /DOMESTIC FINANCE Esc bln Portugal /PRTCBA-OECD STA |
| 3655231A | | Bank credit to economy /Domestic finance /DOMESTIC FINANCE Esc bln Portugal /PRTCBA-OECD S |
| 3656009H | * | EFFECTIVE EXCHANGE RATES 1990 = 100 Portugal /OECD-OECD STATISTICS, PARIS" |
| 365601AH | | EXCHANGE RATES National currency units per US dollar Portugal /OECD-OECD STATISTICS, PARIS |
| 365611AS | * | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Portugal /INTIM |
| 3656151A | | Net foreign position /Foreign finance /FOREIGN FINANCE Esc bln Portugal /PRTCBA-OECD STATI |
| 36765103 | | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE Esc bln Portugal /PRTNSO-OECD |
| 36765303 | | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE Esc bln Portugal /PRTNSO-OECD STATISTICS, |
| 36765603 | *** | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE Esc bln Portugal /PRTNSO-OECD STATISTICS, |

Spain

| OECD Code | OECD Definition |
|-----------|--|
| 3220349H | Consumer goods /Industrial production /PRODUCTION 1990=100 Spain /ESPNSO-OECD STATISTICS, |
| 3220439H | Intermediate goods /Industrial production /PRODUCTION 1990=100 Spain /ESPNSO-OECD STATISTI |
| 3220449H | Investment goods /Industrial production /PRODUCTION 1990=100 Spain /ESPNSO-OECD STATISTICS |
| 3220459K | *** Manufacturing, sa /Industrial production /PRODUCTION 1990=100 Spain /ESPNSO-OECD STATISTIC |
| 3220519J | ** Total, sa /Industrial production /PRODUCTION 1990=100 Spain /ESPECO-OECD STATISTICS, PARIS |
| 32206180 | Passenger cars /Commodity output /PRODUCTION '000 Spain /ESPCAR-OECD STATISTICS, PARIS" |
| 32206580 | Cement /Commodity output /PRODUCTION tonnes '000 Spain /ESPIND-OECD STATISTICS, PARIS" |
| 32206780 | Crude steel /Commodity output /PRODUCTION tonnes '000 Spain /INTISI-OECD STATISTICS, PARIS |
| 32321580 | Dwellings completed /Construction - General /CONSTRUCTION '000 Spain /ESPTRA-OECD STATISTI |
| 32321780 | ** CONSTRUCTION Thousands; monthly averages Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3232239H | * Building construction /Cost of construction /CONSTRUCTION 1990=100 Spain /ESPTRA-OECD STAT |
| 3232319H | Naval construction /Commodity output /PRODUCTION 1990=100 Spain /ESPNSO-OECD STATISTICS, P |
| 32325383 | New passenger car registrations, sa /Domestic trade - other /DOMESTIC TRADE '000 Spain /ES |
| 32428282 | ** Registered unemployed, sa /Unemployment /LABOUR '000 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 324286A3 | STANDARDISED UNEMPLOYMENT RATES, sa Per cent Spain /INTEUR-OECD STATISTICS, PARIS" |
| 32429780 | Labour disputes: time lost /Labour - other /LABOUR '000 Spain /ESPLAB-OECD STATISTICS, PAR |
| 32429983 | * Unfilled vacancies, sa /Labour - other /LABOUR '000 Spain /ESPEMP-OECD STATISTICS, PARIS" |
| 3243419H | * Agricultural products /Producer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PAR |
| 3243479H | Consumer goods /Producer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3243519H | Energy /Producer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3243569H | * PRODUCER PRICES (manufacturing) 1990 = 100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3243649H | Intermediate goods /Producer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3243659H | Investment goods /Producer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3244449H | Fuel and electricity /Consumer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARI |
| 3244459H | *** Food /Consumer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3244559H | *** All items less food /Consumer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS |
| 3244589H | * Rent /Consumer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3244599H | *** Services less rent /Consumer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3244619H | *** All items /Consumer prices /PRICES 1990=100 Spain /ESPNSO-OECD STATISTICS, PARIS" |
| 3254829D | MONETARY AGGREGATES, sa 1990 = 100 Spain /ESPCBA-OECD STATISTICS, PARIS" |
| 3254832A | ** Monetary aggregate (M3) /Domestic finance - General /DOMESTIC FINANCE Ptas bln Spain /ESPC |
| 3254833D | *** ESP MONETARY AGGREGATE M3 SA /BN PESETA Spain OECD STATISTICS, PARIS" |
| 3254839D | *** MONETARY AGGREGATES, sa 1990 = 100 Spain /ESPCBA-OECD STATISTICS, PARIS" |
| 3254861A | *** Total liquidity (ALP2) /Domestic finance - General /DOMESTIC FINANCE Ptas bln Spain /ESPCB |
| 3255302A | Commercial banks /Credit to private sector /DOMESTIC FINANCE Ptas bln Spain /ESPCBA-OECD S |
| 3255312A | Other credit institutions /Credit to private sector /DOMESTIC FINANCE Ptas bln Spain /ESPC |
| 3255631H | *** Call money /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Spain /ESPCBA-OECD STATIS |
| 325564AH | 3-month interbank loans /Interest rates /INTEREST RATES - SHARE PRICES % p.a. Spain /ESPCB |
| 3256009H | EFFECTIVE EXCHANGE RATES 1990 = 100 Spain /OECD-OECD STATISTICS, PARIS" |
| 325601AH | EXCHANGE RATES National currency units per US dollar Spain /OECD-OECD STATISTICS, PARIS" |
| 325611AS | Official reserves excluding gold /Foreign finance /FOREIGN FINANCE SDR mln Spain /INTIMF-O |
| 3256151A | Net foreign position /Foreign finance /FOREIGN FINANCE Ptas bln Spain /ESPCBA-OECD STATIST |
| 32765102 | Net trade (f.o.b.-c.i.f.), sa /Foreign trade /FOREIGN TRADE Ptas bln Spain /ESPCUS-OECD ST |
| 32765252 | FOREIGN TRADE - Ftr Trade Balance (fob-fob), sa Billions US dollars; monthly averages Spai |
| 32765302 | Imports c.i.f., sa /Foreign trade /FOREIGN TRADE Ptas bln Spain /ESPCUS-OECD STATISTICS, P |
| 32765552 | * FOREIGN TRADE - Ftr Imports (fob/cif) Total, sa Billions US dollars; monthly averages Spai |
| 32765602 | Exports f.o.b., sa /Foreign trade /FOREIGN TRADE Ptas bln Spain /ESPCUS-OECD STATISTICS, P |
| 32765752 | FOREIGN TRADE - Ftr Exports Fob Total, sa Billions US dollars; monthly averages Spain /ESP |